

High-Dimensional Statistics & Sparsity

UDRC Summer School

João F. C. Mota

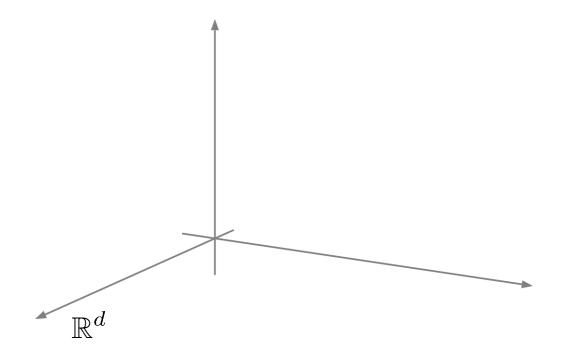
Heriot-Watt University, Edinburgh, UK



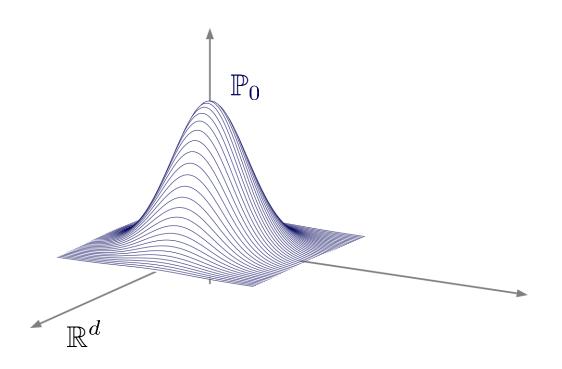




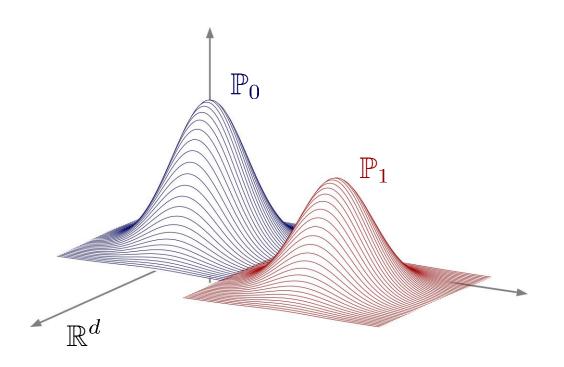






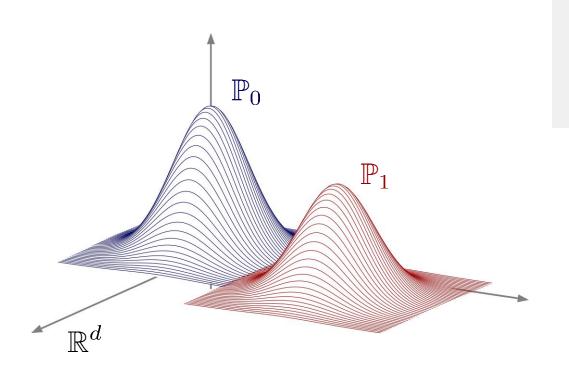








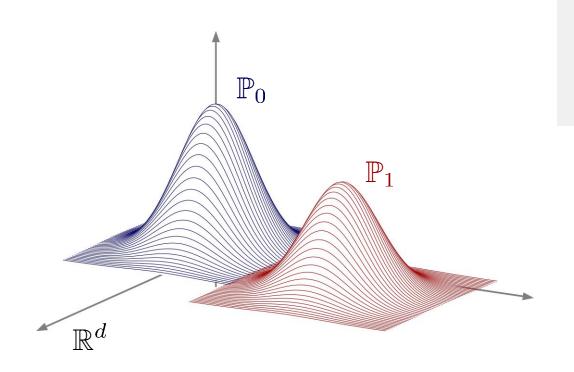
Hypothesis testing in high-dimensions



Problem



Hypothesis testing in high-dimensions

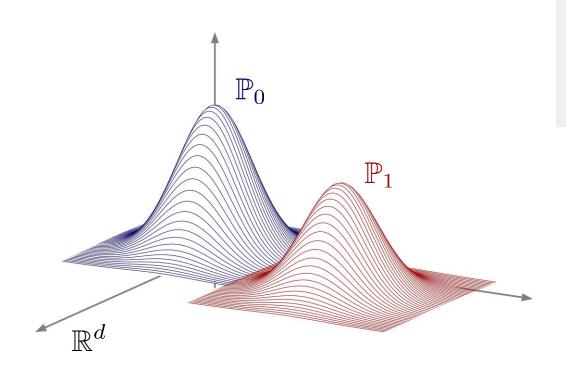


Problem

Observe a random vector $X \in \mathbb{R}^d$



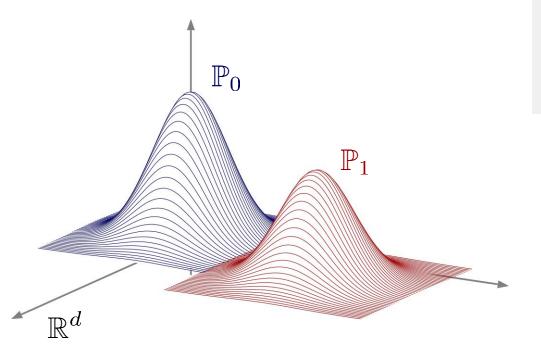
Hypothesis testing in high-dimensions



ProblemObserve a random vector $X \in \mathbb{R}^d$ $X \sim \mathbb{P}_0$ or $X \sim \mathbb{P}_1$?



Hypothesis testing in high-dimensions

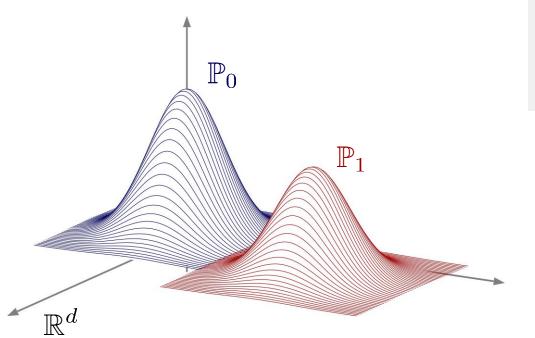


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False positive: $\alpha = \mathbb{P}(\text{decide } X \sim \mathbb{P}_1 \mid X \sim \mathbb{P}_0)$



Hypothesis testing in high-dimensions



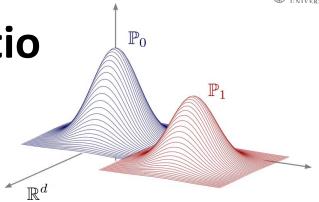
Problem
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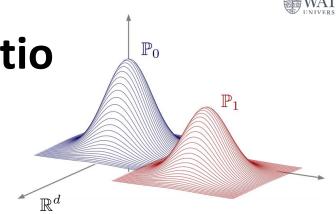
False negative: $\beta = \mathbb{P}(\text{decide } X \sim \mathbb{P}_0 \mid X \sim \mathbb{P}_1)$







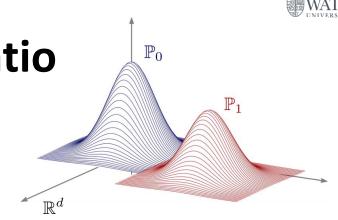
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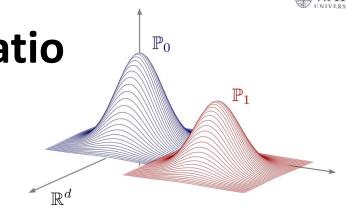




 x_1, \ldots, x_n : i.i.d. realizations of X

For a given $\ T\geq 0$,

If
$$\frac{\mathbb{P}_1(x_1,\ldots,x_n)}{\mathbb{P}_0(x_1,\ldots,x_n)} > T$$



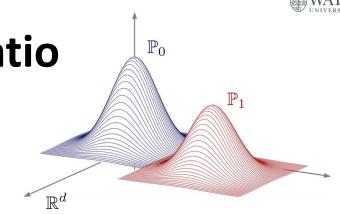


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 \mathbb{T} , then decide $X\sim \mathbb{P}_1$





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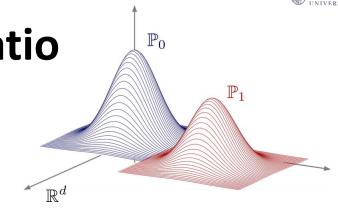
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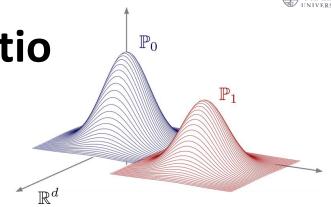
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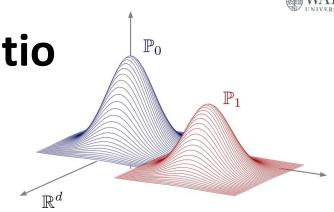
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False negative: β_L

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tio R^d



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$$\mathbf{R}^{d}$$

A

MAP rule / minimizes risk when $T = \frac{\mathbb{P}(X \sim \mathbb{P}_0)}{\mathbb{P}(X \sim \mathbb{P}_1)}$

If
$$\frac{\mathbb{P}_1(x_1,\ldots,x_n)}{\mathbb{P}_0(x_1,\ldots,x_n)} \le T$$

, then decide
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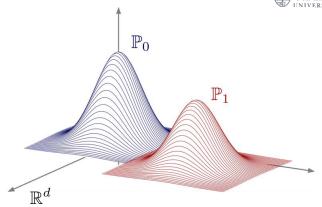
positive:
$$\alpha_L = \mathbb{P}(L(x_1, \dots, x_n) > T \mid X \sim \mathbb{P}_0)$$

TTD (

False p

False negative: $\beta_L = \mathbb{P}(L(x_1, \dots, x_n) \leq T \mid X \sim \mathbb{P}_1)$

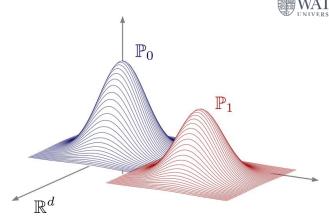






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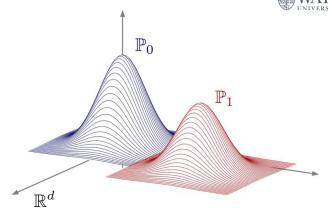
$$\beta_L = \mathbb{P}(L(x_1, \dots, x_n) \le T \mid X \sim \mathbb{P}_1)$$





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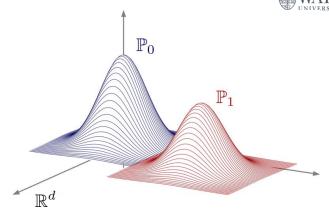


Neyman-Pearson Lemma



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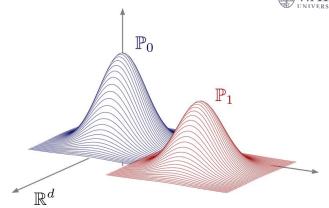
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The likelihood ratio test is *optimal*:



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Neyman-Pearson Lemma

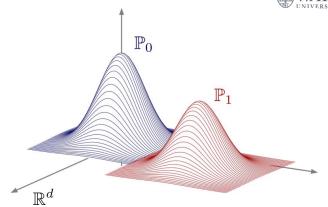
The likelihood ratio test is *optimal*:

If there is another (possibly random) decision rule $D(x_1, \ldots, x_n)$ such that



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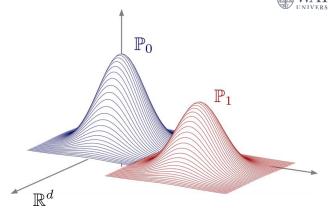
If there is another (possibly random) decision rule $D(x_1, \ldots, x_n)$ such that

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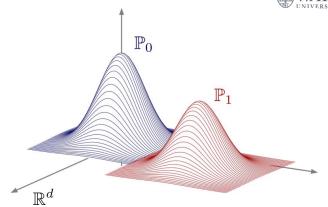
then

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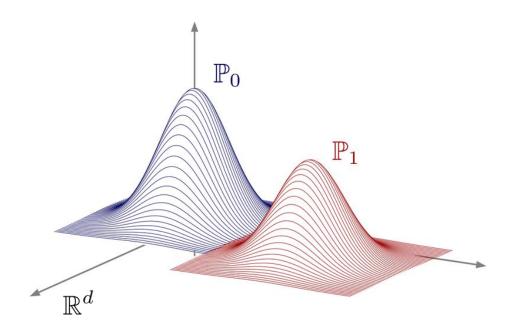
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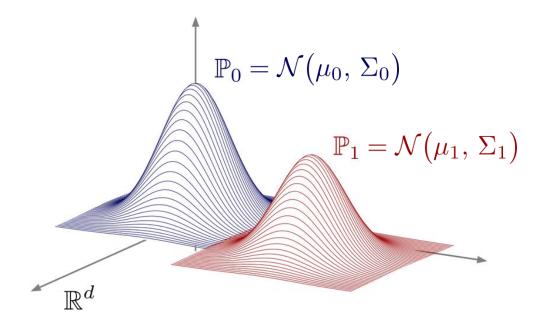
And vice-versa.



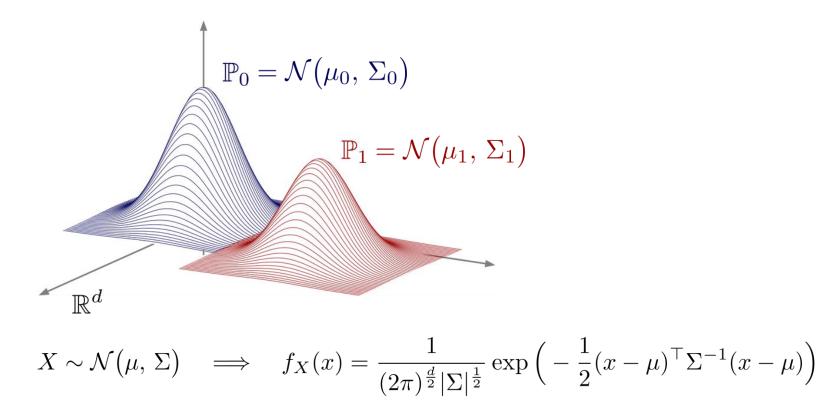














$$\mathbb{P}_{0} = \mathcal{N}(\mu_{0}, \Sigma_{0})$$

$$\mathbb{P}_{1} = \mathcal{N}(\mu_{1}, \Sigma_{1})$$

$$\mathbb{R}^{d}$$

$$X \sim \mathcal{N}(\mu, \Sigma) \implies f_{X}(x) = \frac{1}{(2\pi)^{\frac{d}{2}} |\Sigma|^{\frac{1}{2}}} \exp\left(-\frac{1}{2}(x-\mu)^{\top} \Sigma^{-1}(x-\mu)\right)$$

Assume (to simplify): $\Sigma_0 = \Sigma_1 = \Sigma$



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L(x) > T



Linear Discriminant Analysis

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Assume (to simplify): $\Sigma_0 = \Sigma_1 = \Sigma$ and n = 1 (one observation)

$$L(x) > T \quad \Longleftrightarrow \quad \Psi(x) := \left(x - \frac{\mu_0 + \mu_1}{2}\right)^\top \Sigma^{-1} \left(\mu_1 - \mu_0\right) > \log T$$



$$\Psi(x) := \left(x - \frac{\mu_0 + \mu_1}{2}\right)^\top \Sigma^{-1} \left(\mu_1 - \mu_0\right) > \log T$$



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 $\operatorname{Err}(\Psi) = \mathbb{P}(\Psi(X) > 0 \& \mathbb{P}_0 \operatorname{true}) + \mathbb{P}(\Psi(X) \le 0 \& \mathbb{P}_1 \operatorname{true})$



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(using Gaussianity and manipulating...)



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classical error expression



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| | |
need to be estimated



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 n_0 and n_1 samples



$$\Psi(x) := \left(x - \frac{\mu_0 + \mu_1}{2}\right)^\top \Sigma^{-1} \left(\mu_1 - \mu_0\right) > \log T$$

$$\operatorname{Err}(\Psi) = \mathbb{P}(\Psi(X) > 0 \& \mathbb{P}_0 \operatorname{true}) + \mathbb{P}(\Psi(X) \le 0 \& \mathbb{P}_1 \operatorname{true})$$
$$= \mathbb{P}(\Psi(X) > 0 \mid \mathbb{P}_0) \cdot \mathbb{P}(\mathbb{P}_0) + \mathbb{P}(\Psi(X) \le 0 \mid \mathbb{P}_1) \cdot \mathbb{P}(\mathbb{P}_1)$$
$$= \frac{1}{2} \mathbb{P}(\Psi(X) > 0 \mid \mathbb{P}_0) + \frac{1}{2} \mathbb{P}(\Psi(X) \le 0 \mid \mathbb{P}_1)$$

(using Gaussianity and manipulating...)

$$= \Phi\left(-\frac{\gamma}{2}\right) := \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{-\gamma/2} e^{-t^2/2} dt$$

classical error expression

$$\begin{split} \gamma = \sqrt{(\mu_0 - \mu_1)^\top \Sigma^{-1} (\mu_0 - \mu_1)} \\ \big| \quad \big| \quad \big| \\ \text{need to be estimated} \end{split}$$

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<u>High-dimensional regime</u>: n_0 and n_1 same order as d





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Fisher linear discriminant function





Assume $\mathbb{P}(\mathbb{P}_0) = \mathbb{P}(\mathbb{P}_1)$ $\Sigma = I_d$ d = 400 $n_0 = n_1 = 800$



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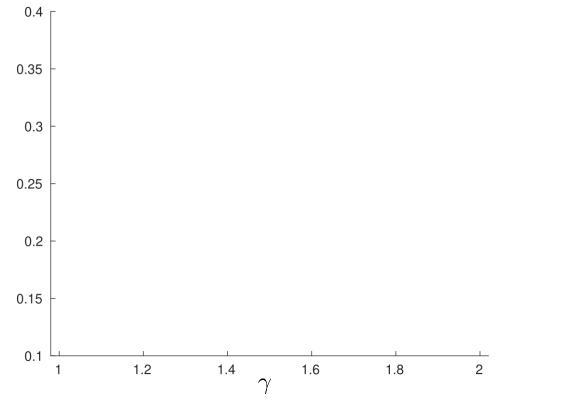
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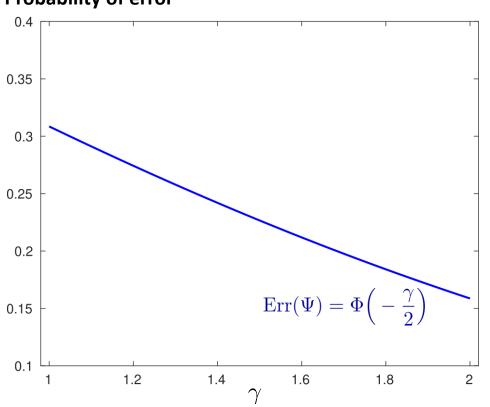
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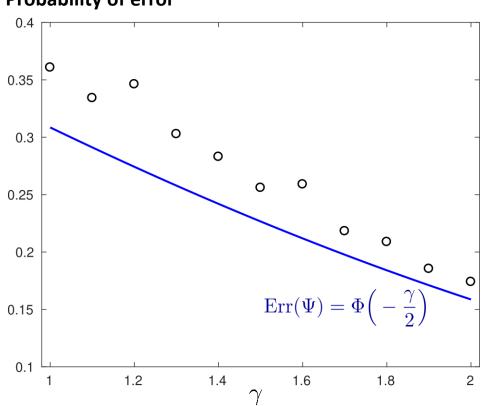
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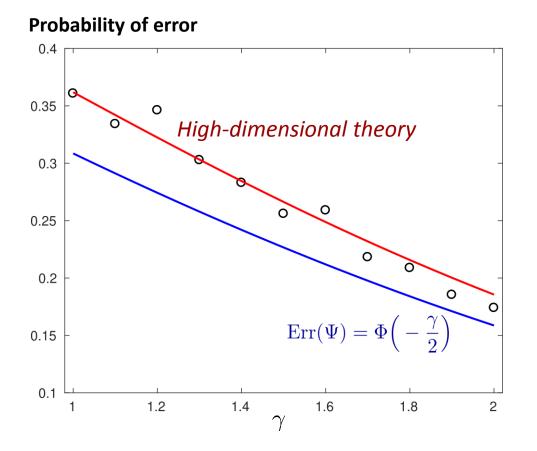




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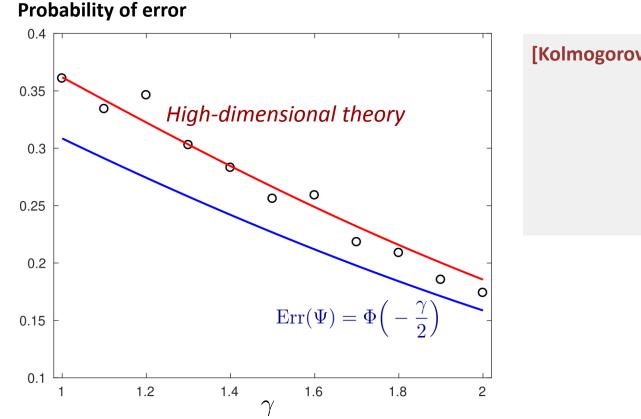




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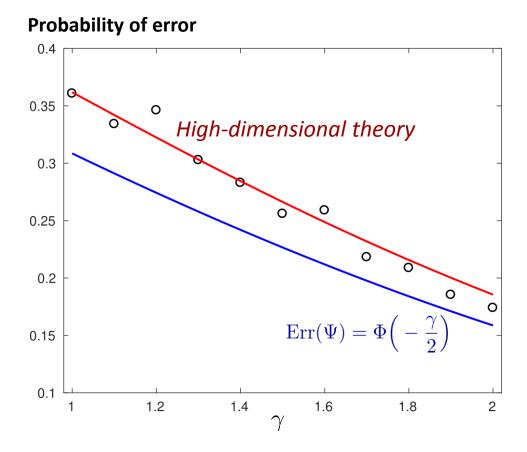


[Kolmogorov]		



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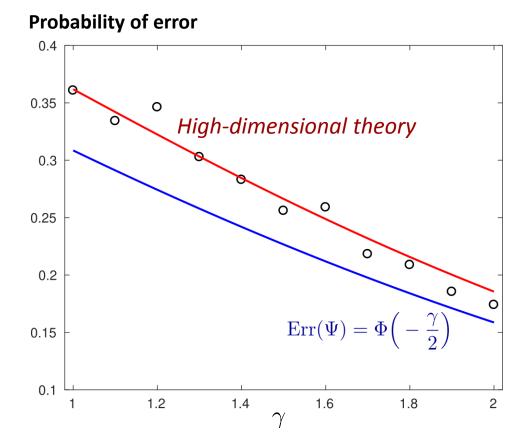


[Kolmogorov]	
$(d, n_0, n_1) \to \infty$	$\frac{d}{n_0}, \ \frac{d}{n_1} \to \alpha$



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What can help in high-dimensions?



What can help in high-dimensions?

Structure



What can help in high-dimensions?

Structure e.g., <u>sparsity</u>



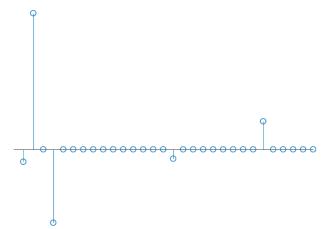
Structure e.g., <u>sparsity</u>

Suppose μ_0 and μ_1 are sparse: only have $s \ll d$ nonzero entries



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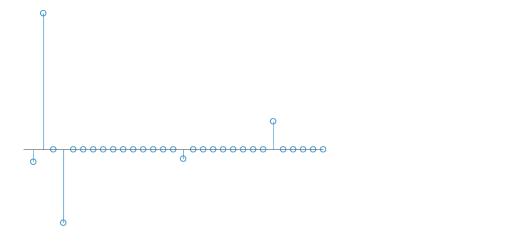
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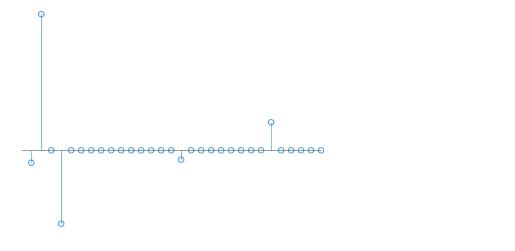


Procedure: *hard-threshold* entries of estimates



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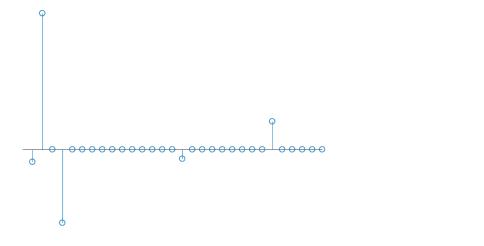
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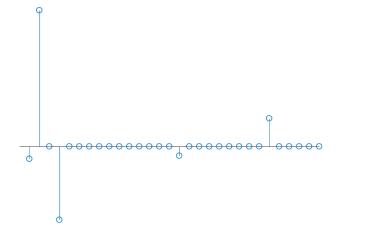
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$$\widehat{\mu}_0' = \frac{1}{n_0} \sum_{i=1}^{n_0} x_i \quad \longrightarrow \quad (\widehat{\mu}_0)_i = H_\lambda \left(\left(\widehat{\mu}_0' \right)_i \right)$$



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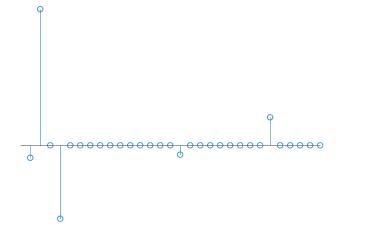
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(same for μ_1)

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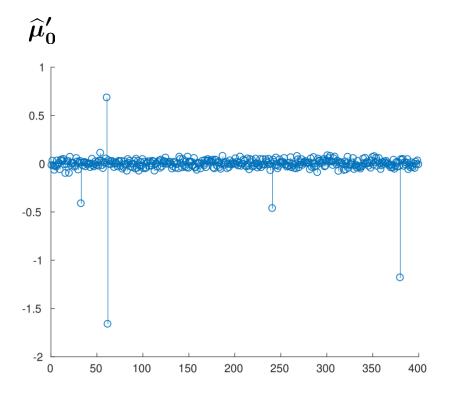
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 $n = 800$ $s = 5$ $\lambda = \sqrt{\frac{2\log d}{n}} = 0.1224$



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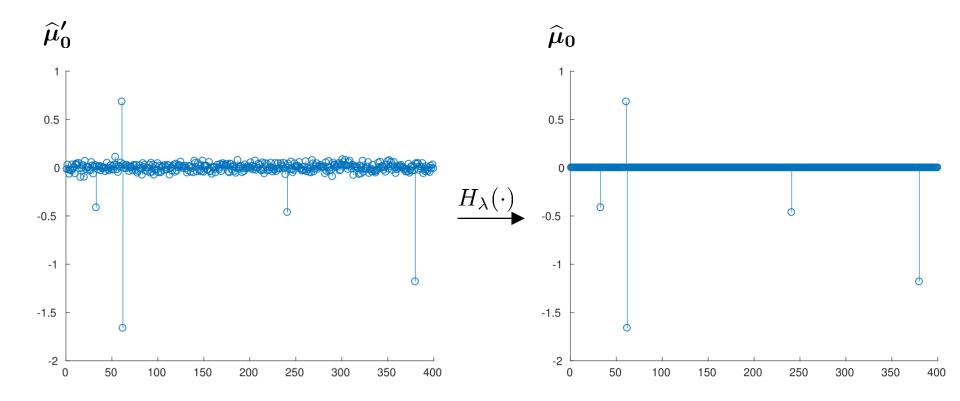
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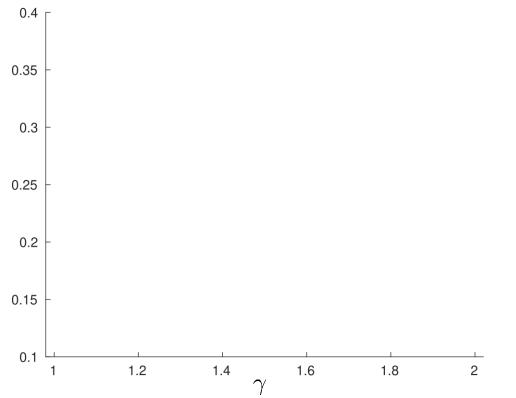
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Probability of error



$$s = 5$$
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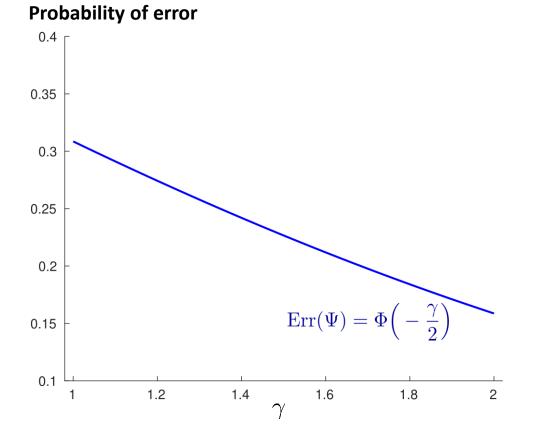


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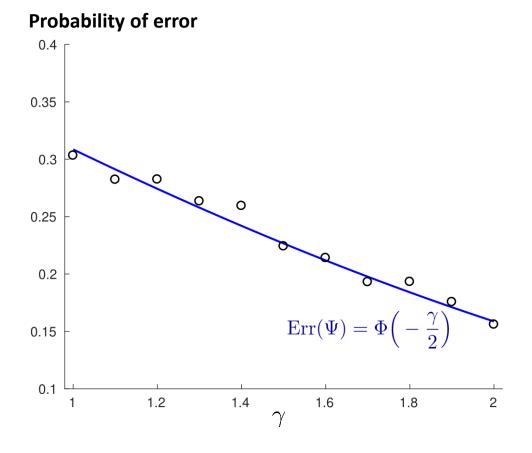




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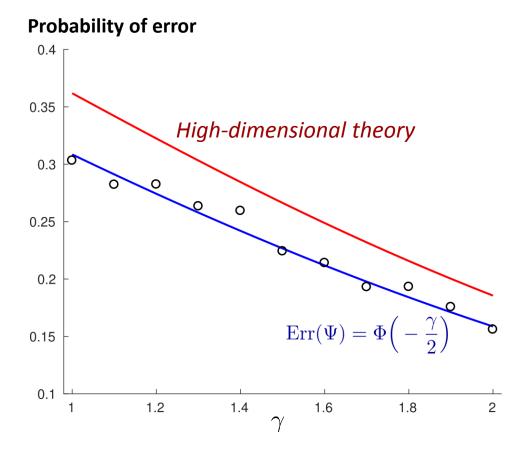
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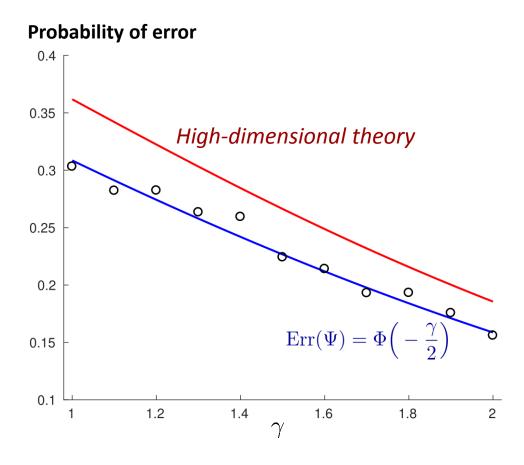
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$$s = 5$$
 $\lambda = \sqrt{\frac{2\log d}{n}}$



Sparsity makes problem low-dimensional



Outline

Motivation: Hypothesis Testing in High-Dimensions

Introduction to LASSO and other sparsity problems

Gaussian graphical model selection

Matrix completion



A Crime Problem

93/40



A Crime Problem

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2	32	72	11	43	18	494
3	57	70	18	16	16	643
4	31	71	11	25	19	341
5	67	72	9	29	24	773
6	25	68	8	32	15	603
7	34	68	12	24	14	484
8	33	62	13	28	11	546
9	36	69	7	25	12	424
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50	66	67	26	18	16	940



A Crime Problem

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50	66	67	26	18	16	940

<u>Goal:</u> Predict *# crimes / million* based on the other indicators



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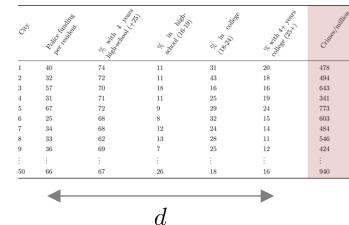
Linear/affine model

ic.	Q ^O CO ^{CO} CO	ale high and high ale	1 Contraction of the second se	64 (1) 20 (1) 40 10 (2) (1) (2) (1) (1) (1) (1) (1) (1) (1) (1) (1) (1	and the second s	Crime willion
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Linear/affine model

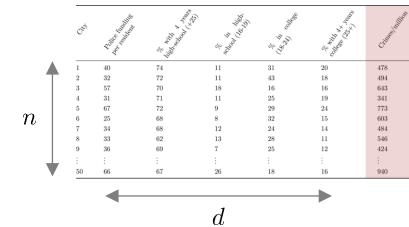
d=5 predictors





Linear/affine model

- d=5 predictors
- n=50 samples

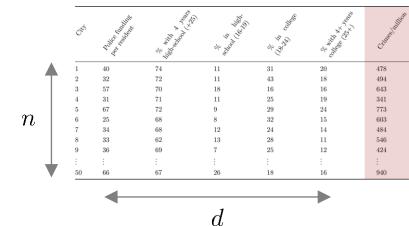




Linear/affine model

d=5 predictors

$$y_i \simeq x_0 + \sum_{j=1}^d a_{ij} x_j \quad i = 1, \dots, n$$





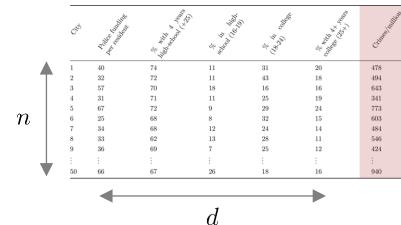
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d=5 predictors

n=50 samples

$$y_i \simeq x_0 + \sum_{j=1}^d a_{ij} x_j \quad i = 1, \dots, n$$

response variable (crime rate)





Linear/affine model

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$$| offset$$

$$response variable$$

$$(crime rate)$$



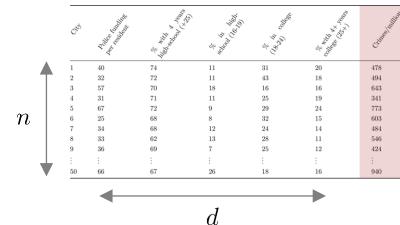


Linear/affine model

d=5 predictors

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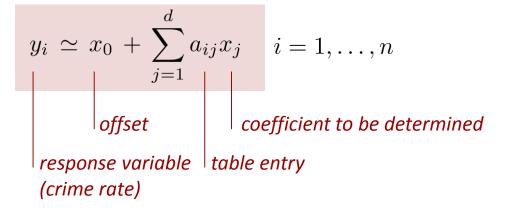
$$| offset \\ response variable \\ (crime rate) \qquad table entry$$

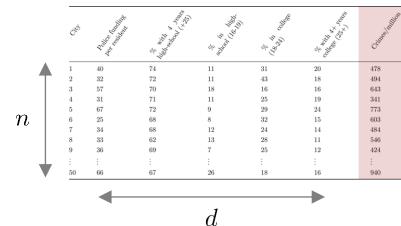




Linear/affine model

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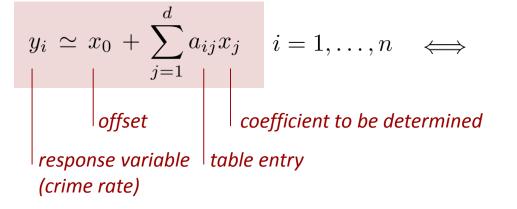


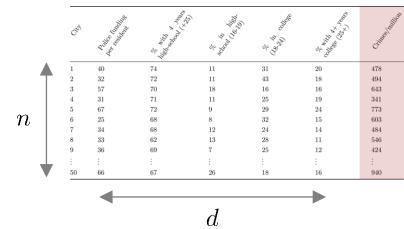


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d = 5 predictors

n = 50 samples



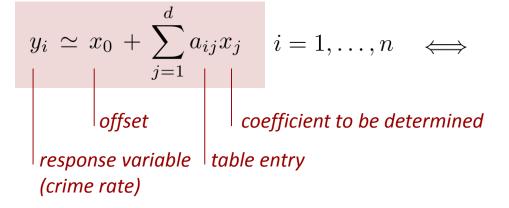


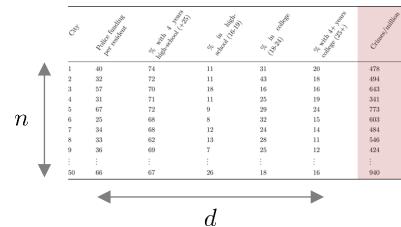
 $y \simeq x_0 1_n + Ax$



Linear/affine model

d = 5 predictors



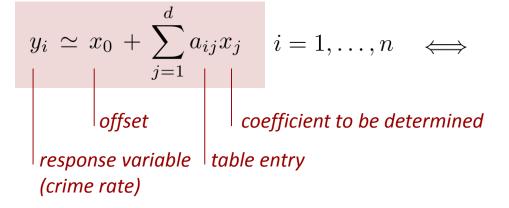


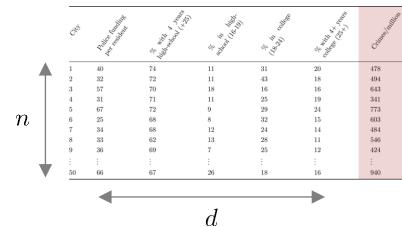
y	\simeq	$x_0 1_n$	+	Ax
	=	$\left[1_n\right]$	A]	$\begin{bmatrix} x_0 \\ x \end{bmatrix}$

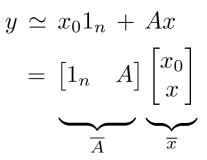


Linear/affine model

d = 5 predictors





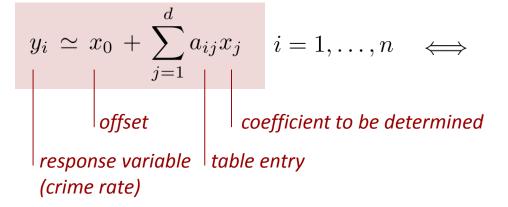




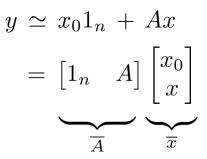
Linear/affine model

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nd



Find coefficients:



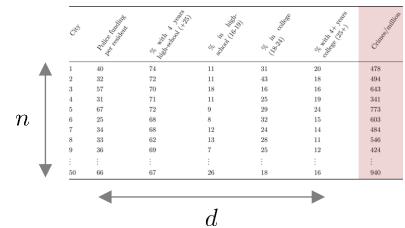
Linear/affine model

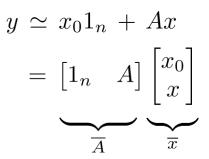
d = 5 predictors

n=50 samples

$$y_{i} \simeq x_{0} + \sum_{j=1}^{d} a_{ij}x_{j} \qquad i = 1, \dots, n \quad \iff$$

$$| offset | coefficient to be determined table entry (crime rate)$$





Find coefficients: *least-squares*

$$\underset{\overline{x}}{\operatorname{minimize}} \quad \frac{1}{2} \left\| y - \overline{A} \overline{x} \right\|_2^2$$



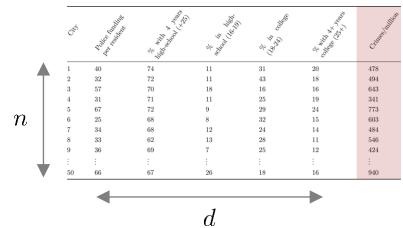
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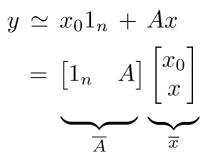
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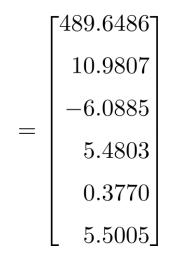
$$\underset{\overline{x}}{\operatorname{minimize}} \quad \frac{1}{2} \left\| y - \overline{A} \overline{x} \right\|_2^2$$

$$\overline{x}_{\rm LS}^{\star} = \left(\overline{A}^{\top}\overline{A}\right)^{-1}\overline{A}\,y$$



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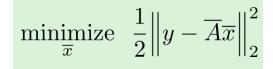


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	489.6486]	offset
	10.9807	
_	-6.0885	
=	5.4803	
	0.3770	
	5.5005	

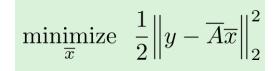




$\overline{x}_{\rm LS}^{\star} = \left(\overline{A}^{\top}\overline{A}\right)^{-1}\overline{A}\,y$

	489.6486]	offset
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Problems with least-squares



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Problems with least-squares

little interpretability



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Problems with least-squares

little interpretability

all coefficients contribute to prediction



$$\underset{\overline{x}}{\operatorname{minimize}} \quad \frac{1}{2} \left\| y - \overline{A} \overline{x} \right\|_2^2$$

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Problems with least-squares

little interpretability

all coefficients contribute to prediction

small bias, large variance



121/40

Linear Regression

$$\underset{\overline{x}}{\operatorname{minimize}} \quad \frac{1}{2} \left\| y - \overline{A} \overline{x} \right\|_2^2$$

$$\overline{x}_{\rm LS}^{\star} = \left(\overline{A}^{\top}\overline{A}\right)^{-1}\overline{A}\,y$$

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Problems with least-squares

little interpretabilityall coefficients contribute to predictionsmall bias, large variancezeroing coefficients can improve mean-squared error





least absolute selection and shrikange operator



least absolute selection and shrikange operator

minimize $\frac{1}{2n} \|y - Ax\|_{2}^{2} + \lambda \|x\|_{1}$



least absolute selection and shrikange operator



least absolute selection and shrikange operator

$$\begin{array}{c|c} \underset{x}{\text{minimize}} & \frac{1}{2n} \|y - Ax\|_{2}^{2} + \lambda \|x\|_{1} - \underbrace{\text{L1-norm}}_{\text{regularization parameter}} \geq 0 \end{array}$$



least absolute selection and shrikange operator

$$\begin{array}{c|c} \underset{x}{\text{minimize}} & \frac{1}{2n} \|y - Ax\|_{2}^{2} + \lambda \|x\|_{1} - \dots \\ & & \|x\|_{1} = |x_{1}| + |x_{2}| + \dots + |x_{d}| \\ & & \|x\|_{1} = |x_{1}| + |x_{2}| + \dots + |x_{d}| \\ & & \|x\|_{1} = |x_{1}| + |x_{2}| + \dots + |x_{d}| \\ & & \|x\|_{1} = |x_{1}| + |x_{2}| + \dots + |x_{d}| \\ & & \|x\|_{1} = |x_{1}| + |x_{2}| + \dots + |x_{d}| \\ & & \|x\|_{1} = |x_{1}| + |x_{2}| + \dots + |x_{d}| \\ & & \|x\|_{1} = |x_{1}| + |x_{2}| + \dots + |x_{d}| \\ & & \|x\|_{1} = |x_{1}| + |x_{2}| + \dots + |x_{d}| \\ & & \|x\|_{1} = |x_{1}| + |x_{2}| + \dots + |x_{d}| \\ & \|x\|_{1} = |x_{1}| + |x_{2}| + \dots + |x_{d}| \\ & \|x\|_{1} = |x_{1}| + |x_{2}| + \dots + |x_{d}| \\ & \|x\|_{1} = |x_{1}| + \|x_{2}\|_{1} + \dots + \|x_{d}\|_{1} \\ & \|x\|_{1} = |x_{1}| + \|x_{2}\|_{1} + \dots + \|x_{d}\|_{1} \\ & \|x\|_{1} = \|x_{1}\|_{1} + \|x_{2}\|_{1} + \dots + \|x_{d}\|_{1} \\ & \|x\|_{1} = \|x_{1}\|_{1} + \|x_{2}\|_{1} + \dots + \|x_{d}\|_{1} \\ & \|x\|_{1} = \|x_{1}\|_{1} + \|x_{2}\|_{1} + \dots + \|x_{d}\|_{1} \\ & \|x\|_{1} = \|x_{1}\|_{1} + \|x_{2}\|_{1} + \dots + \|x_{d}\|_{1} \\ & \|x\|_{1} = \|x_{1}\|_{1} + \|x_{2}\|_{1} + \dots + \|x_{d}\|_{1} \\ & \|x\|_{1} = \|x_{1}\|_{1} + \|x_{2}\|_{1} + \dots + \|x_{d}\|_{1} \\ & \|x\|_{1} = \|x\|_{1} + \|x\|_{1} \\ & \|x\|_{1} = \|x\|_{1} + \|x\|_{1} +$$



least absolute selection and shrikange operator

$$\begin{array}{cccc} \underset{x}{\text{minimize}} & \frac{1}{2n} \|y - Ax\|_{2}^{2} + \lambda \|x\|_{1} - \dots & \text{L1-norm} & \|x\|_{1} = |x_{1}| + |x_{2}| + \dots + |x_{d}| \\ & & \\ &$$

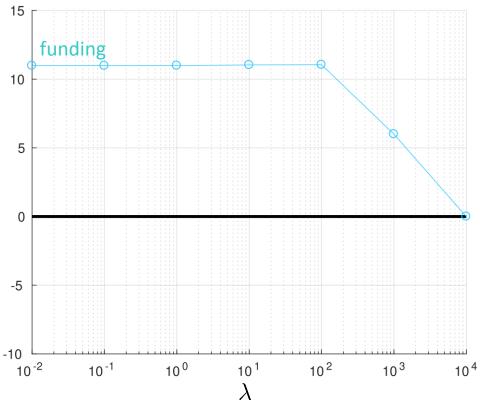
Coefficient value

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least absolute selection and shrikange operator

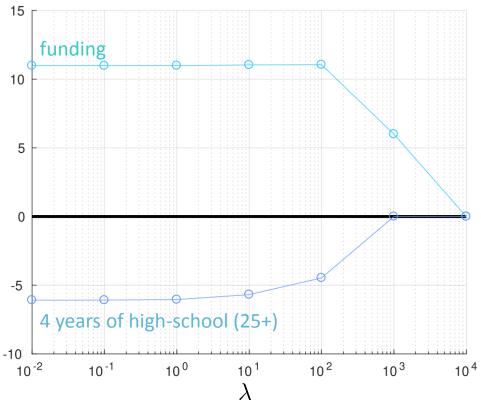
Coefficient value





least absolute selection and shrikange operator

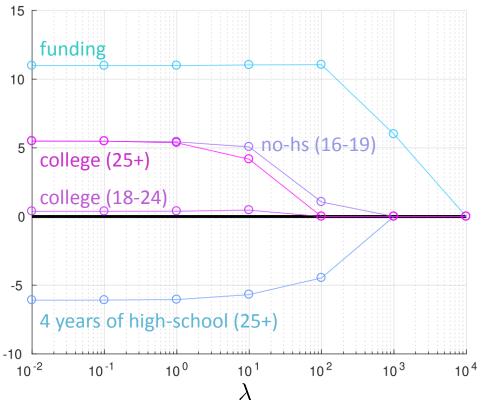
Coefficient value





least absolute selection and shrikange operator



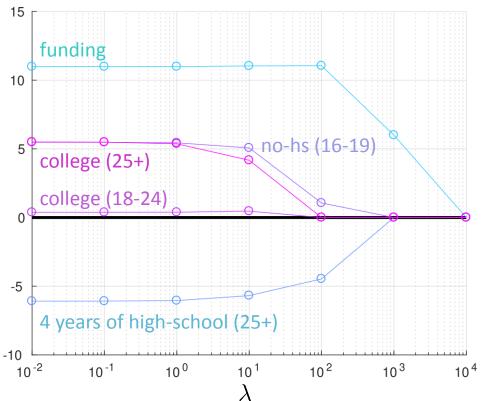




least absolute selection and shrikange operator

$$\begin{array}{cccc} \underset{x}{\text{minimize}} & \frac{1}{2n} \|y - Ax\|_{2}^{2} + \lambda \|x\|_{1} - \frac{L1 - norm}{2} \|x\|_{1} = |x_{1}| + |x_{2}| + \dots + |x_{d}| \\ & & \\ &$$

Coefficient value

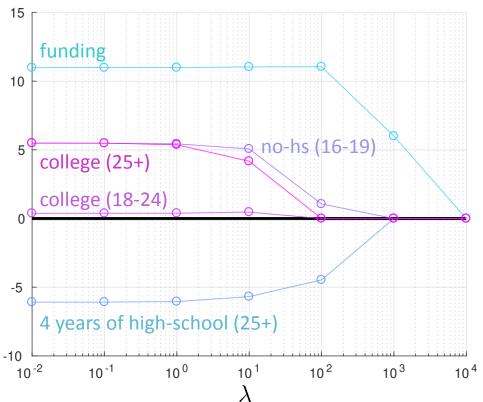


In reality, we solved ...



least absolute selection and shrikange operator

Coefficient value



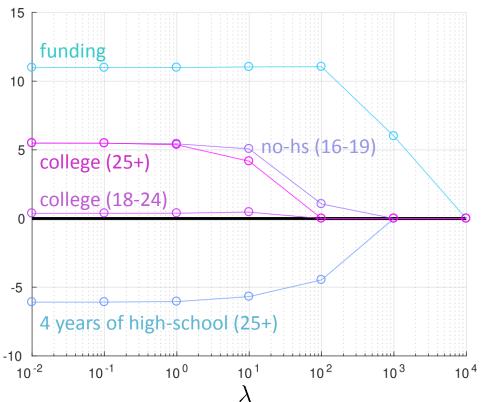
In reality, we solved ...

$$\min_{x_0, x} \frac{1}{2n} \|y - x_0 \mathbf{1}_n - Ax\|_2^2 + \lambda \|x\|_1$$



least absolute selection and shrikange operator

Coefficient value



In reality, we solved ...

$$\min_{x_0, x} \frac{1}{2n} \|y - x_0 \mathbf{1}_n - Ax\|_2^2 + \lambda \|x\|_1$$

$$\| \\ \| \\ \text{necessary because } \frac{1}{n} \sum_{i=1}^n y_i \neq 0$$





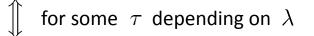
minimize
$$\frac{1}{2n} \|y - Ax\|_2^2 + \lambda \|x\|_1$$

LASSO (aka Basis Pursuit Denoising)



minimize
$$\frac{1}{2n} \|y - Ax\|_{2}^{2} + \lambda \|x\|_{1}$$

LASSO (aka Basis Pursuit Denoising)





minimize
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LASSO (aka Basis Pursuit Denoising)

$$igfluon$$
 for some $\, au\,$ depending on $\,\lambda\,$

$$\begin{array}{ll} \underset{x}{\text{minimize}} & \left\| y - Ax \right\|_{2}^{2} \\ \text{subject to} & \|x\|_{1} \leq \tau \end{array}$$

Constrained LASSO



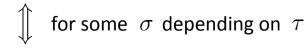
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Constrained LASSO





minimize
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LASSO (aka Basis Pursuit Denoising)

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$$\ensuremath{\Uparrow}$$
 for some $\,\sigma\,$ depending on $\,\tau\,$

$$\begin{array}{ll} \underset{x}{\text{minimize}} & \|x\|_{1} \\ \text{subject to} & \left\|y - Ax\right\|_{2} \leq \sigma \end{array}$$

Constrained LASSO

Relaxed Basis Pursuit



minimize
$$\frac{1}{2n} \|y - Ax\|_{2}^{2} + \lambda \|x\|_{1}$$

LASSO (aka Basis Pursuit Denoising)

$$igcap_{}$$
 for some $\, au\,$ depending on $\,\lambda$

$$\begin{array}{ll} \underset{x}{\text{minimize}} & \left\| y - Ax \right\|_{2}^{2} \\ \text{subject to} & \|x\|_{1} \leq \tau \end{array}$$

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$$\begin{array}{ll} \underset{x}{\text{minimize}} & \|x\|_{1} \\ \text{subject to} & \left\|y - Ax\right\|_{2} \leq \sigma \end{array}$$

Constrained LASSO

Relaxed Basis Pursuit

Basis Pursuit when $\sigma = 0$



 $\begin{aligned} \widehat{x} \in & \underset{x}{\operatorname{arg\,min}} & \|x\|_{1} \\ & \text{s.t.} & \|y - Ax\|_{2} \leq \sigma \end{aligned}$



Assume $\sigma = 0$:



$\begin{aligned} \widehat{x} \in & \underset{x}{\operatorname{arg\,min}} & \|x\|_{1} \\ & \text{s.t.} & \|y - Ax\|_{2} \leq \sigma \end{aligned}$

Assume $\sigma = 0$:

y = Ax has solutions



Assume $\sigma = 0$:

y = Ax has solutions $\widetilde{x} + \operatorname{null}(A)$



$$\widehat{x} \in \underset{x}{\operatorname{arg\,min}} \|x\|_{1} \\ \text{s.t.} \|y - Ax\|_{2} \leq \sigma$$

$$y = Ax$$
 has solutions $\widetilde{x} + \operatorname{null}(A)$
$$y = A\widetilde{x} \Big|$$



$$\widehat{x} \in \underset{x}{\operatorname{arg\,min}} \|x\|_{1} \\ \text{s.t.} \|y - Ax\|_{2} \leq \sigma$$

$$y = Ax$$
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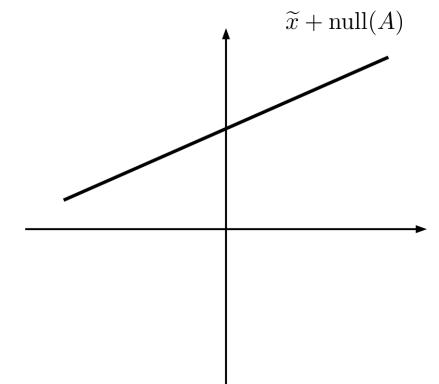
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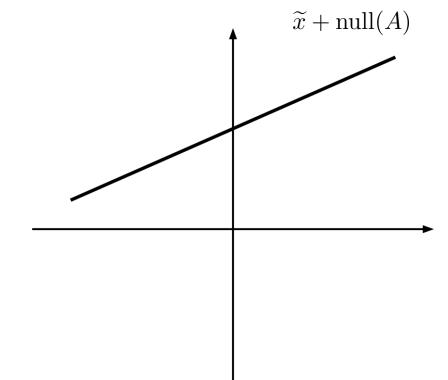




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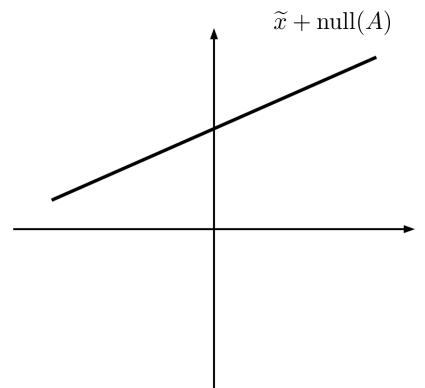
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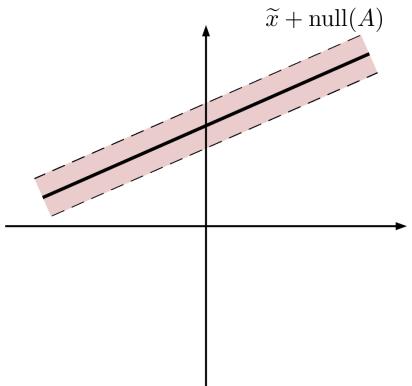
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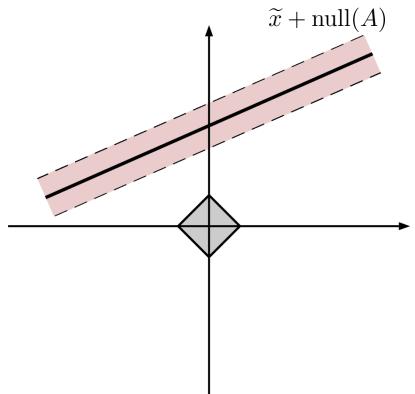




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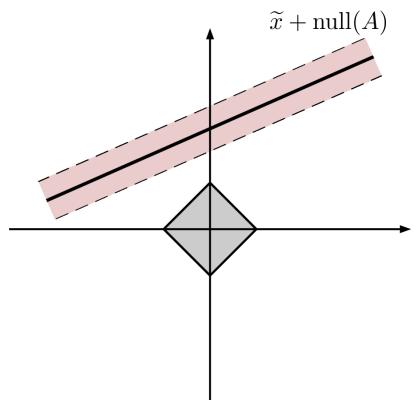




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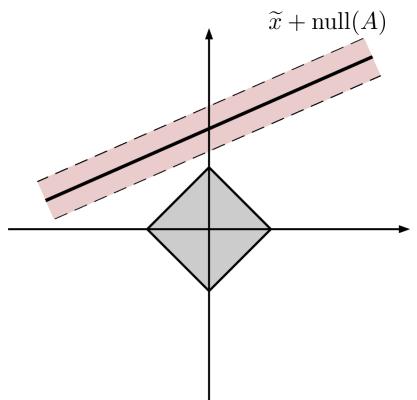




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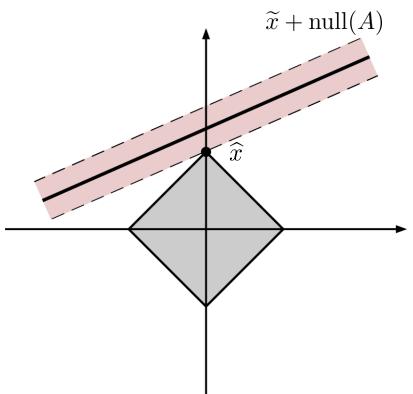




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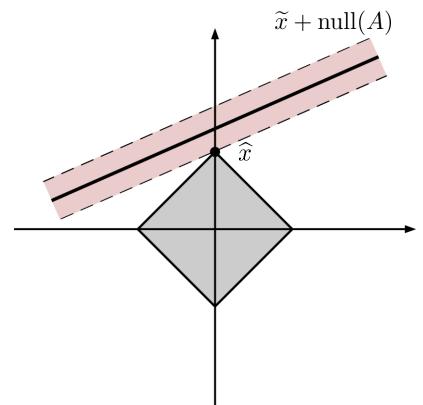
Assume $\sigma = 0$:

$$y = Ax$$
 has solutions $\widetilde{x} + \operatorname{null}(A)$
 $y = A\widetilde{x} \Big| \quad \Big| \{d : Ad = 0\}$

Assume $\sigma > 0$: margin around $\tilde{x} + \operatorname{null}(A)$

$$\widehat{x}_2 \in \underset{x}{\operatorname{arg\,min}} \|x\|_2^2$$

s.t. $\|y - Ax\|_2 \leq \sigma$





$$\widehat{x} \in \underset{x}{\operatorname{arg\,min}} \|x\|_{1} \\ \text{s.t.} \|y - Ax\|_{2} \le \sigma$$

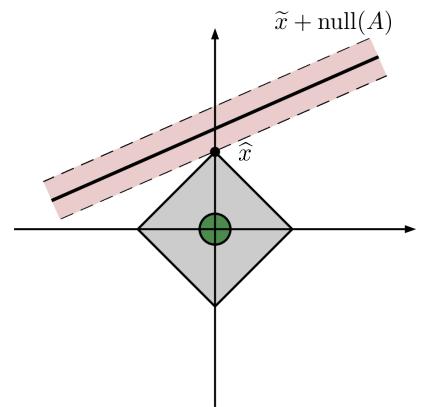
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 $y = A\widetilde{x} \Big| \quad \Big| \{d : Ad = 0\}$

Assume $\sigma > 0$: margin around $\tilde{x} + \operatorname{null}(A)$

$$\widehat{x}_2 \in \underset{x}{\operatorname{arg\,min}} \|x\|_2^2$$

s.t. $\|y - Ax\|_2 \le \sigma$





$$\widehat{x} \in \underset{x}{\operatorname{arg\,min}} \|x\|_{1} \\ \text{s.t.} \|y - Ax\|_{2} \le \sigma$$

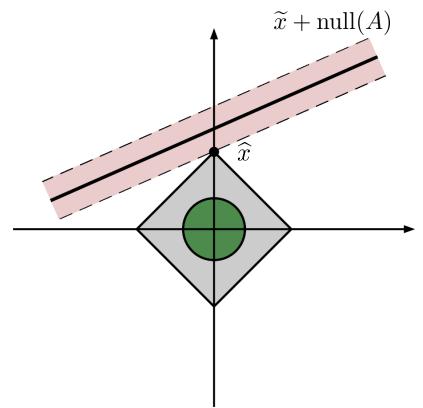
Assume $\sigma = 0$:

$$y = Ax$$
 has solutions $\widetilde{x} + \operatorname{null}(A)$
 $y = A\widetilde{x} \Big| \quad \Big| \{d : Ad = 0\}$

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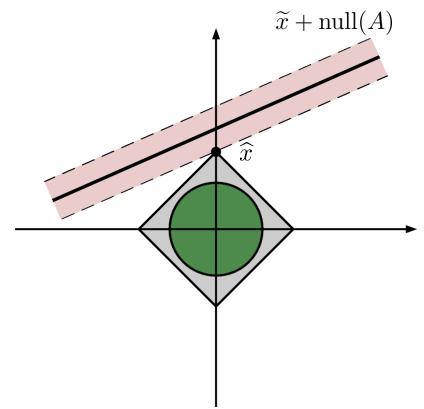
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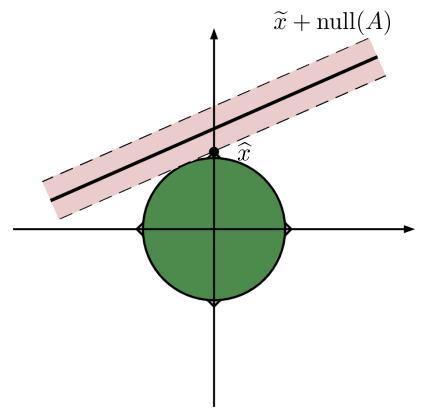
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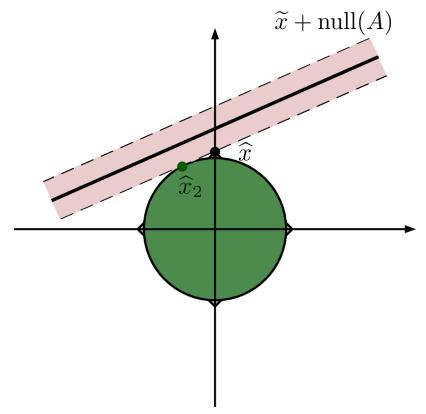
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Assume $\sigma > 0$: margin around $\tilde{x} + \operatorname{null}(A)$

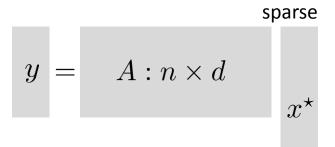
$$\widehat{x}_2 \in \underset{x}{\operatorname{arg\,min}} \|x\|_2^2$$

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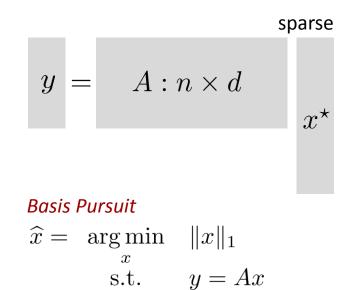






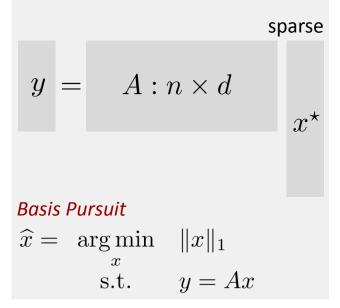






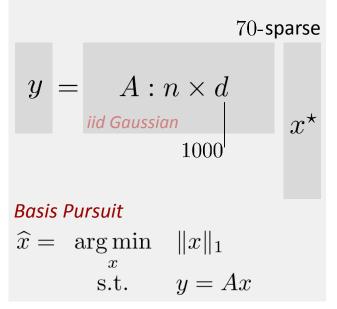


Compressed Sensing (CS)

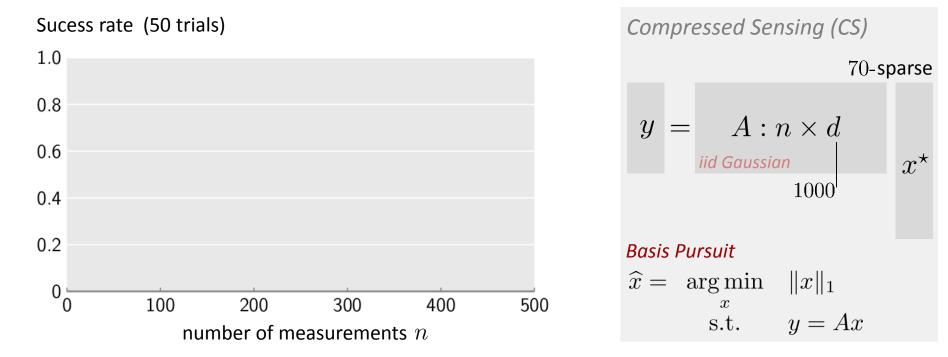




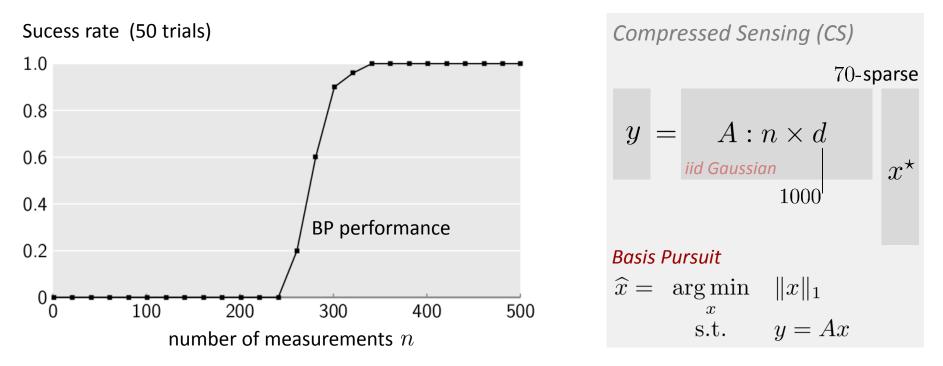
Compressed Sensing (CS)







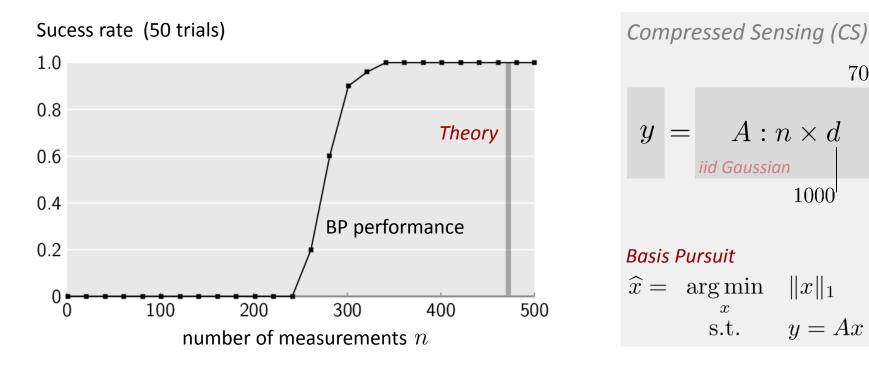






70-sparse

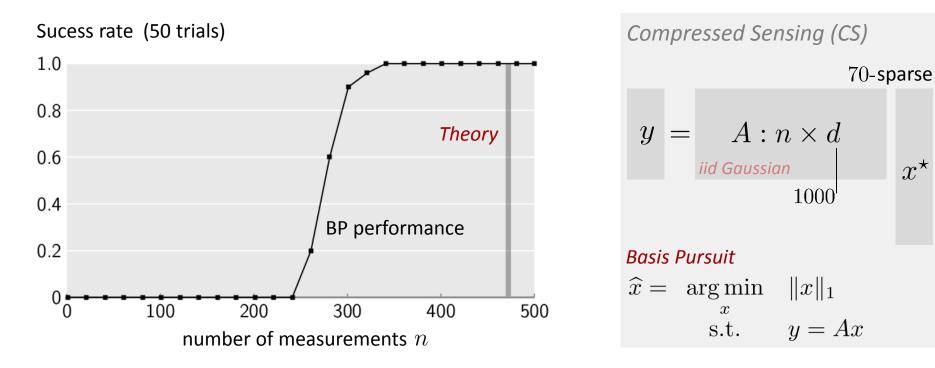
 x^{\star}





 x^{\star}

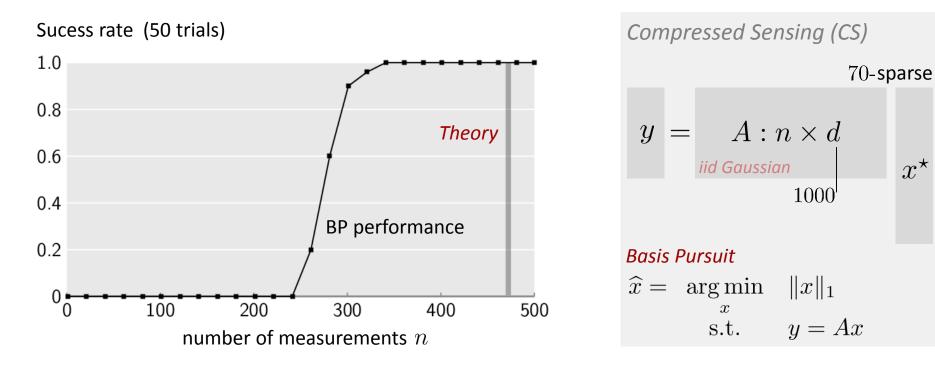
Example: Compressed Sensing



Theorem [Chandrasekaran et al. 12']



 x^{\star}



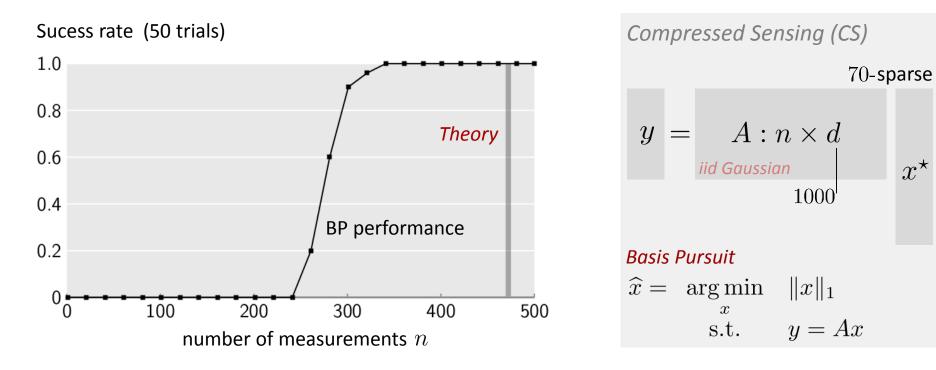
Theorem [Chandrasekaran et al. 12']

 $x^{\star} \in \mathbb{R}^d$



 x^{\star}

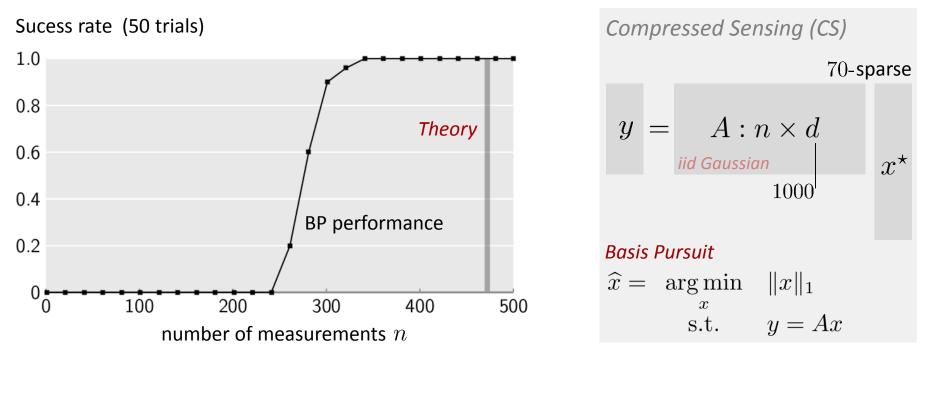
Example: Compressed Sensing



Theorem [Chandrasekaran et al. 12']

 $x^{\star} \in \mathbb{R}^{d}$ unknown, but *s*-sparse



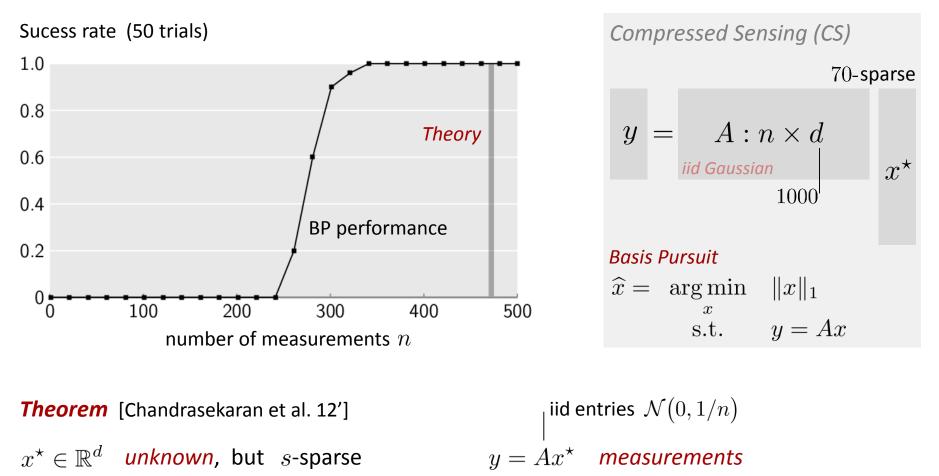


Theorem [Chandrasekaran et al. 12']

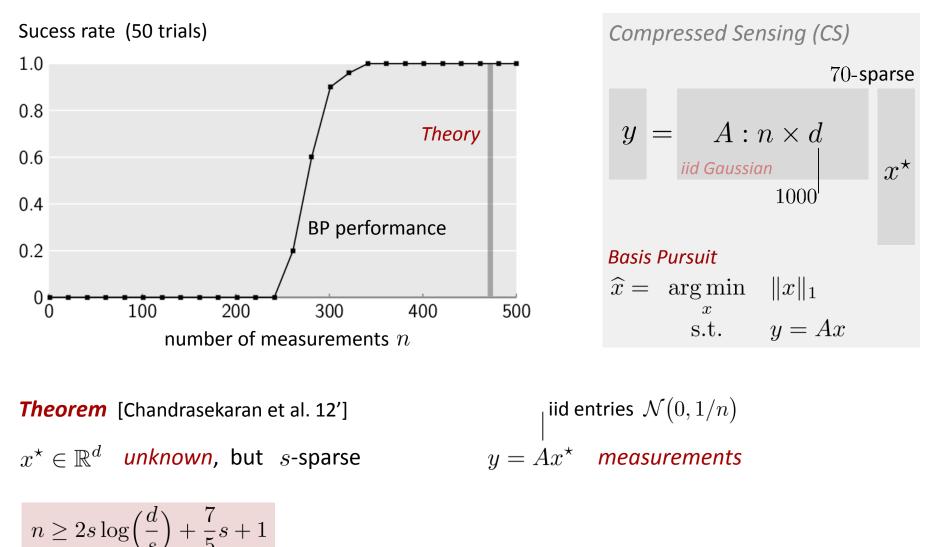
 $x^{\star} \in \mathbb{R}^{d}$ unknown, but *s*-sparse

 $y = Ax^{\star}$ measurements

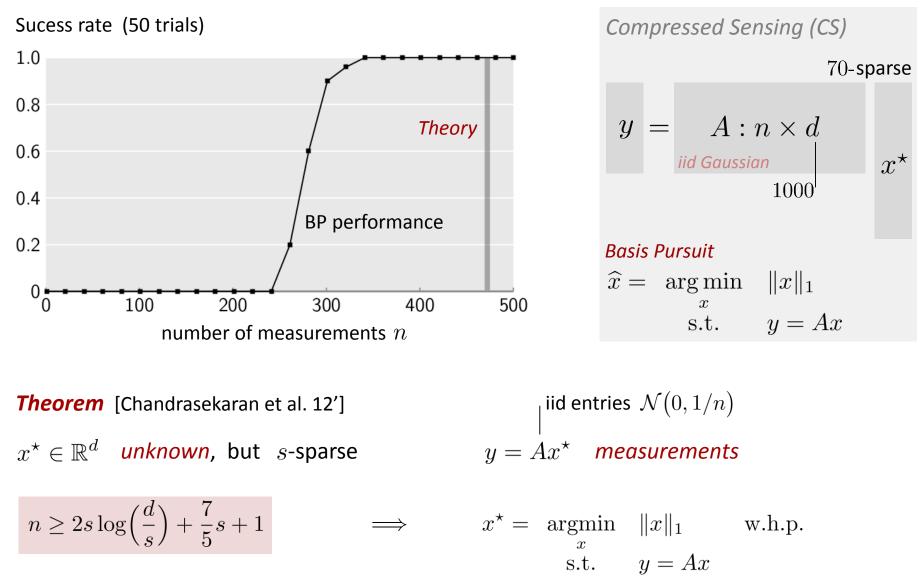














Application: Image Reconstruction



Application: Image Reconstruction





Application: Image Reconstruction



 256×496





 $256 \times 496 \qquad \Longrightarrow \quad z^{\star} \in \mathbb{R}^{126976}$





 $256 \times 496 \qquad \Longrightarrow \quad z^{\star} \in \mathbb{R}^{126976}$

not sparse





 $256 \times 496 \qquad \implies \quad z^{\star} \in \mathbb{R}^{126976}$

not sparse

Natural images have sparse representations





 $256 \times 496 \qquad \Longrightarrow \quad z^{\star} \in \mathbb{R}^{126976}$

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Natural images have sparse representations

$$z^{\star} = \Psi x^{\star}$$





 $256 \times 496 \qquad \implies \quad z^{\star} \in \mathbb{R}^{126976}$

not sparse

Natural images have sparse representations

$$z^{\star} = \Psi x^{\star}$$

sparse or near-sparse





 $256 \times 496 \qquad \implies \quad z^{\star} \in \mathbb{R}^{126976}$

not sparse

Natural images have sparse representations

$$z^{\star} = \Psi x^{\star}$$

sparse or near-sparse *dictionary* (wavelet, DCT, gradient space)





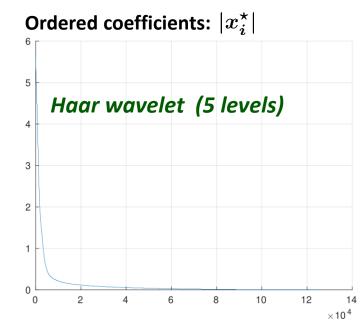
 $256 \times 496 \qquad \Longrightarrow \quad z^{\star} \in \mathbb{R}^{126976}$

not sparse

Natural images have sparse representations

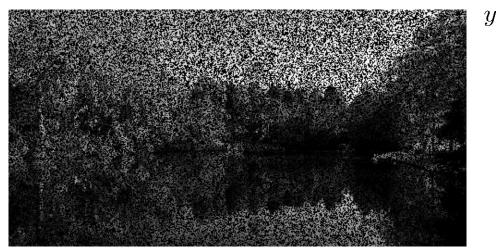
$$z^{\star} = \Psi x^{\star}$$

sparse or near-sparse *dictionary* (wavelet, DCT, gradient space)











Suppose we observe *only 50%* of pixels



Solve $\widehat{x} = \underset{x}{\operatorname{arg\,min}} \|x\|_{1}$ s.t. $y = \Phi \Psi x$





Solve

$$\widehat{x} = \underset{x}{\operatorname{arg\,min}} \|x\|_{1}$$

s.t. $y = \Phi \Psi x$
|wavelet





Solve

$$\widehat{x} = \underset{x}{\operatorname{arg\,min}} \|x\|_{1} \\ \text{s.t.} \quad y = \Phi \Psi x \\ | \underset{\text{observed indices}}{\| \text{wavelet} \|}$$





Solve $\widehat{x} = \arg \min \|x\|_1$ x $y = \Phi \Psi x$ s.t. wavelet observed indices





Suppose we observe *only 50%* of pixels



Solve $\widehat{x} = \arg \min \|x\|_1$ x $y = \Phi \Psi x$ s.t. wavelet observed indices



PSNR: 21.31 dB



Solve

$$\widehat{x} = \underset{x}{\operatorname{arg\,min}} \|x\|_{1}$$

s.t. $y = \Phi \Psi x$
 $\|wavelet$
partial DFT



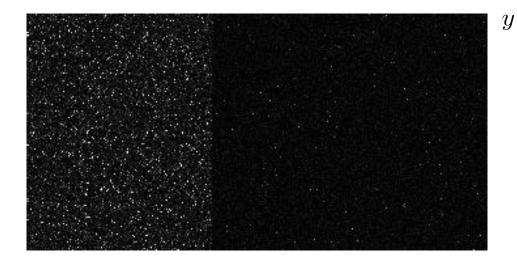
Solve

$$\widehat{x} = \underset{x}{\operatorname{arg\,min}} \|x\|_{1}$$

s.t. $y = \Phi \Psi x$
 $\|wavelet$
partial DFT

each entry of y has info from entire image





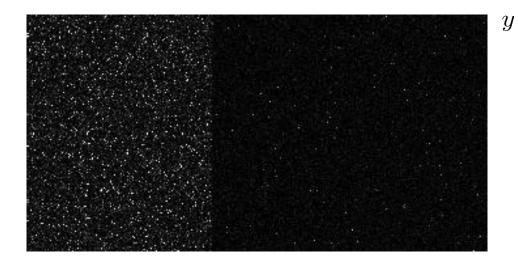
Solve

$$\widehat{x} = \underset{x}{\operatorname{arg\,min}} \|x\|_{1}$$

s.t. $y = \Phi \Psi x$
 $\|wavelet$
partial DFT

each entry of y has info from entire image





Solve $\widehat{x} = \arg \min \|x\|_1$ xs.t. $y = \Phi \Psi x$ wavelet partial DFT

each entry of y has info from entire image



 \widehat{x}

PSNR: 24.93 dB



Outline

Motivation: Hypothesis Testing in High-Dimensions

Introduction to LASSO and other sparsity problems

Gaussian graphical model selection

Matrix completion

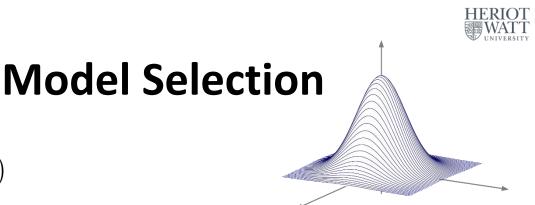




 $X = (X_1, X_2, \dots, X_d)$



 $X = (X_1, X_2, \dots, X_d) \sim \mathcal{N}(0_d, \Sigma^{\star})$



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<u>Problem</u>: given *n* idd observations of *X*, denoted $X^{(1)}, \ldots, X^{(n)}$, estimate Σ^{\star}

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<u>Problem</u>: given *n* idd observations of *X*, denoted $X^{(1)}, \ldots, X^{(n)}$, estimate Σ^{\star}

Assumption:

ERIOT

 $X = (X_1, X_2, \dots, X_d) \sim \mathcal{N}(0_d, \Sigma^{\star})$

<u>Problem</u>: given *n* idd observations of *X*, denoted $X^{(1)}, \ldots, X^{(n)}$, estimate Σ^{\star}

Assumption: most pairs of coordinates (X_i, X_j) are conditionally independent

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<u>Problem</u>: given *n* idd observations of *X*, denoted $X^{(1)}, \ldots, X^{(n)}$, estimate Σ^{\star}

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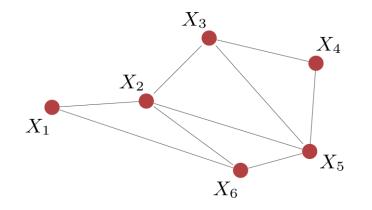
 $\stackrel{\Psi}{\text{precision matrix }} \Theta^{\star} := \left(\Sigma^{\star}\right)^{-1}$ is *sparse*

 $X = (X_1, X_2, \dots, X_d) \sim \mathcal{N}(0_d, \Sigma^{\star})$

<u>Problem</u>: given *n* idd observations of *X*, denoted $X^{(1)}, \ldots, X^{(n)}$, estimate Σ^{\star}

Assumption: most pairs of coordinates (X_i, X_j) are conditionally independent

$$\Downarrow$$
 precision matrix $\Theta^{\star} := (\Sigma^{\star})^{-1}$ is *sparse*

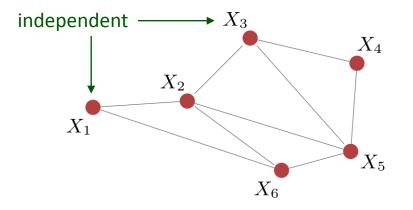


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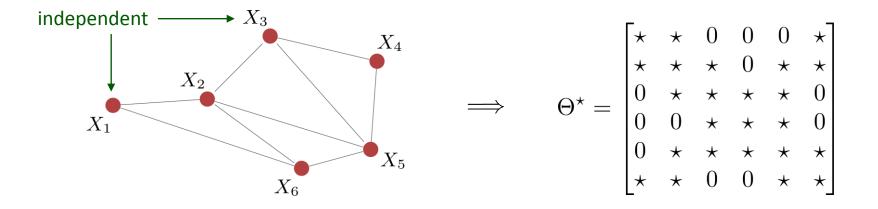


 $X = (X_1, X_2, \dots, X_d) \sim \mathcal{N}(0_d, \Sigma^{\star})$

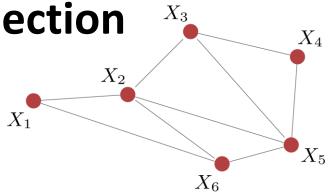
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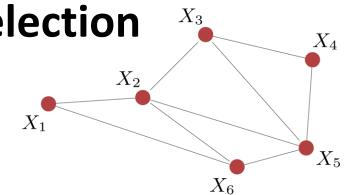








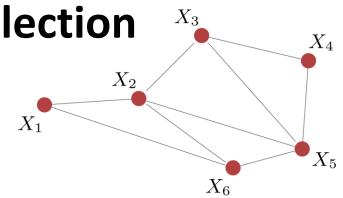
 $X = (X_1, X_2, \dots, X_d) \sim \mathcal{N}(0_d, \Sigma^\star)$





$$X = (X_1, X_2, \dots, X_d) \sim \mathcal{N}(0_d, \Sigma^{\star})$$

$$\mathsf{pdf}: \quad f_X(x;\,\Theta^\star) = \frac{\sqrt{\det\,\Theta^\star}}{(2\pi)^{\frac{d}{2}}} \exp\left(-\frac{1}{2}x^\top\Theta^\star x\right)$$





 X_4

 X_5

 X_3

 X_6

 X_2

 X_1

Gaussian Graphical Model Selection

$$X = (X_1, X_2, \dots, X_d) \sim \mathcal{N}(0_d, \Sigma^{\star})$$

pdf:
$$f_X(x; \Theta^*) = \frac{\sqrt{\det \Theta^*}}{(2\pi)^{\frac{d}{2}}} \exp\left(-\frac{1}{2}x^\top \Theta^* x\right)$$

Maximum likelihood estimator of $\, \Theta^{\star} \,$



 X_4

 X_5

 X_3

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Gaussian Graphical Model Selection

$$X = (X_1, X_2, \dots, X_d) \sim \mathcal{N}(0_d, \Sigma^{\star})$$

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$$\widehat{\Theta}_{\mathrm{ML}} = \underset{\Theta}{\mathrm{arg\,max}} \log \prod_{i=1}^{n} f\left(x^{(1)}, \dots, x^{(n)}; \Theta\right)$$

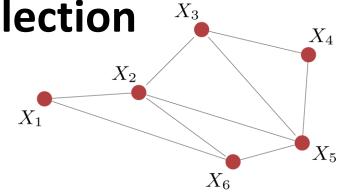


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$$= \underset{\Theta}{\operatorname{arg\,min}} - \log \det \Theta + \operatorname{tr}\left(\Theta \widehat{\Sigma}_{n}\right)$$





 X_4

 X_5

 X_3

 X_6

 X_2

 X_1

Gaussian Graphical Model Selection

$$X = (X_1, X_2, \dots, X_d) \sim \mathcal{N}(0_d, \Sigma^{\star})$$

$$\mathsf{pdf}: \quad f_X(x;\,\Theta^\star) = \frac{\sqrt{\det\,\Theta^\star}}{\left(2\pi\right)^{\frac{d}{2}}} \exp\left(-\frac{1}{2}x^\top\Theta^\star x\right)$$

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$$= \underset{\Theta}{\operatorname{arg\,min}} - \log \det \Theta + \operatorname{tr}\left(\Theta \widehat{\Sigma}_{n}\right)$$
$$|_{sample \ covariance \ matrix}$$





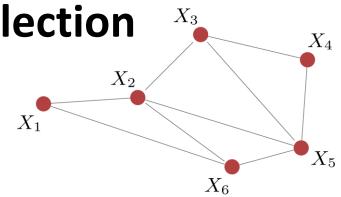
Gaussian Graphical Model Selection

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$$= \underset{\Theta}{\operatorname{arg\,min}} - \log \det \Theta + \operatorname{tr}\left(\Theta \widehat{\Sigma}_{n}\right)$$
$$| sample covariance$$



 $\textit{matrix} \quad \widehat{\Sigma}_n := \frac{1}{n} \sum_{i=1}^n x^{(i)} {x^{(i)}}^\top$

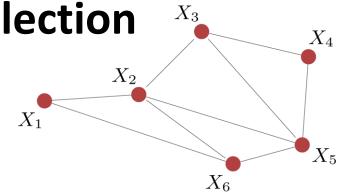


Gaussian Graphical Model Selection

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Maximum likelihood estimator of $\, \Theta^{\star} \,$





 X_4

 X_5

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 X_1

Gaussian Graphical Model Selection

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Maximum likelihood estimator of $\, \Theta^{\star} \,$

(assuming it is invertible \implies n > d)



Gaussian Graphical Model Selection

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$$\mathsf{pdf}: \quad f_X(x;\,\Theta^\star) = \frac{\sqrt{\det\,\Theta^\star}}{(2\pi)^{\frac{d}{2}}} \exp\left(-\frac{1}{2}x^\top\Theta^\star x\right)$$

Maximum likelihood estimator of $\, \Theta^{\star} \,$

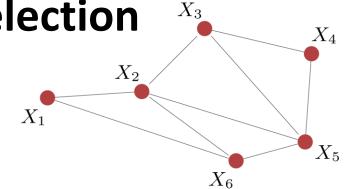
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$$= \underset{\Theta}{\operatorname{arg\,min}} - \log \,\det\,\Theta + \operatorname{tr}\left(\Theta\,\widehat{\Sigma}_n\right) \\ | \\ samp \\ = \widehat{\Sigma}_n^{-1}$$

sample covariance matrix

$$\widehat{\Sigma}_n := \frac{1}{n} \sum_{i=1}^n x^{(i)} x^{(i)^{\top}}$$

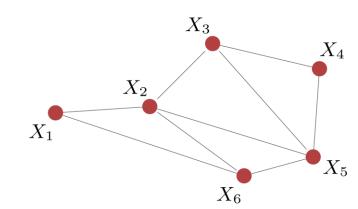
(assuming it is invertible \implies n>d)





Maximum likelihood estimator

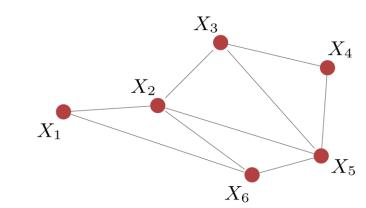
$$\widehat{\Theta}_{\mathrm{ML}}^{\star} = \underset{\Theta}{\operatorname{arg\,min}} - \log \, \det \, \Theta + \operatorname{tr} \left(\Theta \, \widehat{\Sigma}_n \right)$$





Maximum likelihood estimator

$$\widehat{\Theta}_{\mathrm{ML}}^{\star} = \underset{\Theta}{\operatorname{arg\,min}} - \log \, \det \, \Theta + \operatorname{tr} \left(\Theta \, \widehat{\Sigma}_n \right)$$

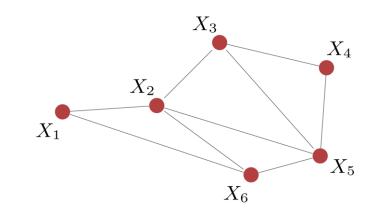


Assumption: most pairs of coordinates (X_i, X_j) are conditionally independent



Maximum likelihood estimator

$$\widehat{\Theta}_{\mathrm{ML}}^{\star} = \underset{\Theta}{\operatorname{arg\,min}} - \log \, \det \, \Theta + \operatorname{tr} \left(\Theta \, \widehat{\Sigma}_n \right)$$



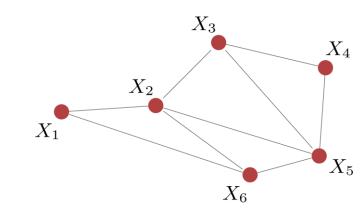
Assumption: most pairs of coordinates (X_i, X_j) are conditionally independent

Graphical LASSO



Maximum likelihood estimator

$$\widehat{\Theta}_{\mathrm{ML}}^{\star} = \underset{\Theta}{\operatorname{arg\,min}} - \log \, \det \, \Theta + \operatorname{tr} \left(\Theta \, \widehat{\Sigma}_n \right)$$



Assumption: most pairs of coordinates (X_i, X_j) are conditionally independent

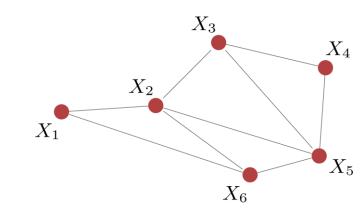
Graphical LASSO

$$\widehat{\Theta}_{\mathrm{GL}} = \underset{\Theta}{\operatorname{arg\,min}} - \log \, \det \, \Theta + \operatorname{tr} \left(\Theta \, \widehat{\Sigma}_n \right) + \lambda \left\| \Theta \right\|_{1, \mathrm{off}\text{-d}}$$



Maximum likelihood estimator

$$\widehat{\Theta}_{\mathrm{ML}}^{\star} = \underset{\Theta}{\operatorname{arg\,min}} - \log \, \det \, \Theta + \operatorname{tr} \left(\Theta \, \widehat{\Sigma}_n \right)$$



Assumption: most pairs of coordinates (X_i, X_j) are conditionally independent

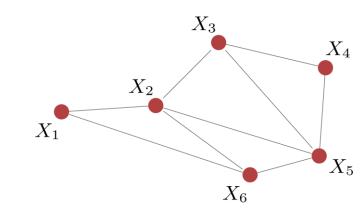
Graphical LASSO

$$\widehat{\Theta}_{\mathrm{GL}} = \underset{\Theta}{\operatorname{arg\,min}} - \log \, \det \, \Theta + \operatorname{tr} \left(\Theta \, \widehat{\Sigma}_n \right) + \frac{\lambda \left\| \Theta \right\|_{1, \mathrm{off}\text{-d}}}{\left| \sum_{i \neq j} \left| \Theta_{ij} \right| \right|}$$



Maximum likelihood estimator

$$\widehat{\Theta}_{\mathrm{ML}}^{\star} = \underset{\Theta}{\operatorname{arg\,min}} - \log \, \det \, \Theta + \operatorname{tr} \left(\Theta \, \widehat{\Sigma}_n \right)$$



Assumption: most pairs of coordinates (X_i, X_j) are conditionally independent

Graphical LASSO

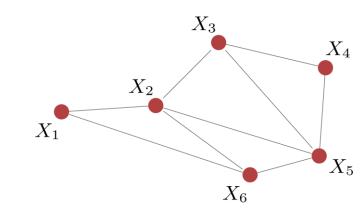
$$\widehat{\Theta}_{\text{GL}} = \underset{\Theta}{\operatorname{arg\,min}} - \log \,\det\,\Theta + \operatorname{tr}\left(\Theta\,\widehat{\Sigma}_n\right) + \frac{\lambda \|\Theta\|_{1,\text{off-d}}}{\left|\sum_{i\neq j} |\Theta_{ij}|\right|}$$

applies L1-norm only to off-diagonal entries



Maximum likelihood estimator

$$\widehat{\Theta}_{\mathrm{ML}}^{\star} = \underset{\Theta}{\operatorname{arg\,min}} - \log \, \det \, \Theta + \operatorname{tr} \left(\Theta \, \widehat{\Sigma}_n \right)$$



Assumption: most pairs of coordinates (X_i, X_j) are conditionally independent

Graphical LASSO

* *

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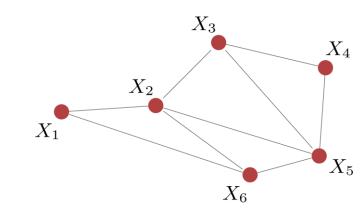
$$\widehat{\Theta}_{\mathrm{GL}} = \underset{\Theta}{\operatorname{arg\,min}} - \log\,\det\,\Theta + \operatorname{tr}\left(\Theta\,\widehat{\Sigma}_{n}\right) + \lambda \|\Theta\|_{1,\mathrm{off-d}}$$
$$\sum_{i\neq j} |\Theta|_{i\neq j} |\Theta|_{i\neq j} = \begin{bmatrix} \star & \star & 0 & 0 & \star \\ \star & \star & 0 & 0 & \star \\ 0 & \star & \star & \star & 0 \\ 0 & 0 & \star & \star & \star & 0 \\ 0 & \star & \star & \star & \star & \star \\ \star & \star & 0 & 0 & \star & \star \end{bmatrix}$$

ijs L1-norm only to off-diagonal entries



Maximum likelihood estimator

$$\widehat{\Theta}_{\mathrm{ML}}^{\star} = \underset{\Theta}{\operatorname{arg\,min}} - \log \, \det \, \Theta + \operatorname{tr} \left(\Theta \, \widehat{\Sigma}_n \right)$$



Assumption: most pairs of coordinates (X_i, X_j) are conditionally independent

Graphical LASSO

$$\widehat{\Theta}_{\mathrm{GL}} = \underset{\Theta}{\operatorname{arg\,min}} - \log\,\det\Theta + \operatorname{tr}\left(\Theta\,\widehat{\Sigma}_n\right) + \lambda \|\Theta\|_{1,\mathrm{off}} - d \\ \left|\sum_{\substack{i\neq j \\ i\neq j}} |\Theta|_{i\neq j} \right| = \left| \begin{array}{c} \star & \star & 0 & 0 & \star \\ \star & \star & \star & 0 & \star & \star \\ 0 & \star & \star & \star & \star & 0 \\ 0 & 0 & \star & \star & \star & \star & 0 \\ 0 & \star & \star & \star & \star & \star & \star \\ \star & \star & 0 & 0 & \star & \star \end{array} \right|$$

 $\left|\sum_{i \neq j} |\Theta_{ij}|\right|$ applies L1-norm only to off-diagonal entries

sensible estimators even for non-Gaussian RVs







Week	Alcoa	$\begin{array}{c} \mathbf{American}\\ \mathbf{Express} \end{array}$	Boeing	Bank of America	Caterpillar	Cisco Systems
1	14.67	43.30	66.15	10.59	100.25	14.94
2	15.29	43.73	69.26	10.89	101.30	15.14
3	15.82	43.86	69.42	11.18	102.59	16.04
4	15.87	43.86	70.29	11.47	102.72	16.41
5	15.92	43.96	70.86	11.87	103.42	16.59
6	15.95	44.13	71.17	11.89	103.56	16.82
7	15.96	44.20	71.43	12.28	104.86	16.88
8	16.18	44.75	71.52	12.32	105.58	16.93
9	16.19	44.94	71.60	12.36	105.87	17.01
÷	:	÷	÷		÷	÷
24	17.42	50.74	79.31	14.77	96.93	21.22
25	18.06	51.39	80.35	15.08	99.62	22.11



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9	16.19	44.94	71.60	12.36	105.87	17.01
:	÷	÷	:	:	÷	:
24	17.42	50.74	79.31	14.77	96.93	21.22
25	18.06	51.39	80.35	15.08	99.62	22.11
	\widetilde{X}_1	\widetilde{X}_2	\widetilde{X}_3	\widetilde{X}_4	\widetilde{X}_5	\widetilde{X}_6



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	\widetilde{X}_1	\widetilde{X}_2	\widetilde{X}_3	\widetilde{X}_4	\widetilde{X}_5	\widetilde{X}_6
remove mean	Ļ	Ļ	Ļ	Ļ	Ļ	Ļ



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	\widetilde{X}_1	\widetilde{X}_2	\widetilde{X}_3	\widetilde{X}_4	\widetilde{X}_5	\widetilde{X}_6
emove mean	Ļ				Ļ	Ļ
	\dot{X}_1	X_2	X_3	X_4	X_5	X_6



Price of stock of *6 companies* at beginning of each week of 2011 (Jan-Jun)

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	\widetilde{X}_1	\widetilde{X}_2	\widetilde{X}_3	\widetilde{X}_4	\widetilde{X}_5	\widetilde{X}_6
remove mean	Ļ	Ļ	Ļ	Ļ	Ļ	Ļ
	X_1	X_2	X_3	X_4	X_5	X_6 η

236/40

25

6





$$\widehat{\Theta}_{\rm ML} = \widehat{\Sigma}_n^{-1} = \begin{bmatrix} 60.82 & 4.85 & -6.21 & -21.34 & -0.07 & -4.81 \\ 4.85 & 7.34 & -1.22 & -5.50 & 0.37 & -4.78 \\ -6.21 & -1.22 & 3.03 & 2.08 & -0.08 & -2.69 \\ -21.34 & -5.50 & 2.08 & 14.31 & -0.42 & 1.69 \\ -0.07 & 0.37 & -0.08 & -0.42 & 0.06 & 0.07 \\ -4.81 & -4.78 & -2.69 & 1.69 & 0.07 & 11.38 \end{bmatrix}$$



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Boeing

Bank of America

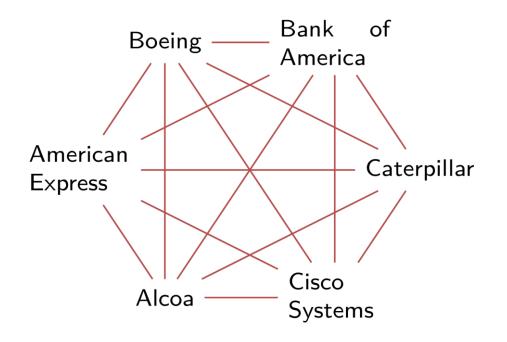
American Express

Caterpillar





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 $\widehat{\Theta}_{\mathrm{GL}} = \underset{\Theta}{\operatorname{arg\,min}} - \log \, \det \, \Theta + \operatorname{tr} \left(\Theta \, \widehat{\Sigma}_n \right) + \lambda \left\| \Theta \right\|_{1, \mathrm{off}\text{-d}}$

graphical LASSO



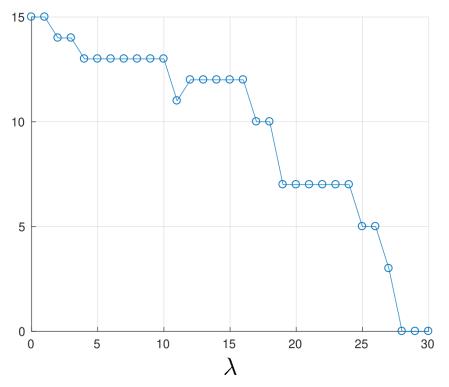
$$\widehat{\Theta}_{\rm GL} = \underset{\Theta}{\operatorname{arg\,min}} - \log \,\det\,\Theta + \operatorname{tr}\left(\Theta\,\widehat{\Sigma}_n\right) + \lambda \|\Theta\|_{1,{\rm off-d}} \quad \text{graphical LASSO}$$
$$\left| (\widehat{\Theta}_{\rm GL})_{ij} \right| \leq 10^{-3} \quad \Longrightarrow \quad \text{we assume no correlation, i.e., no edge } (i,j)$$



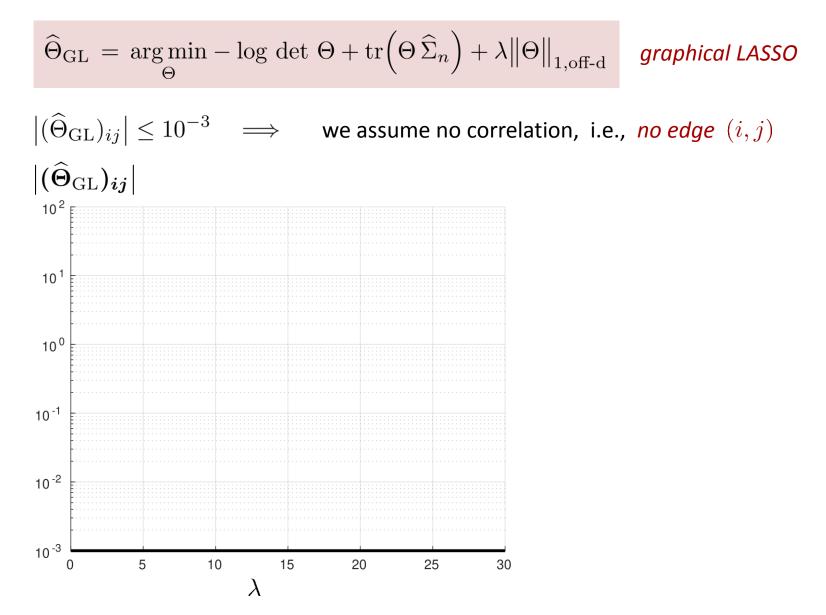
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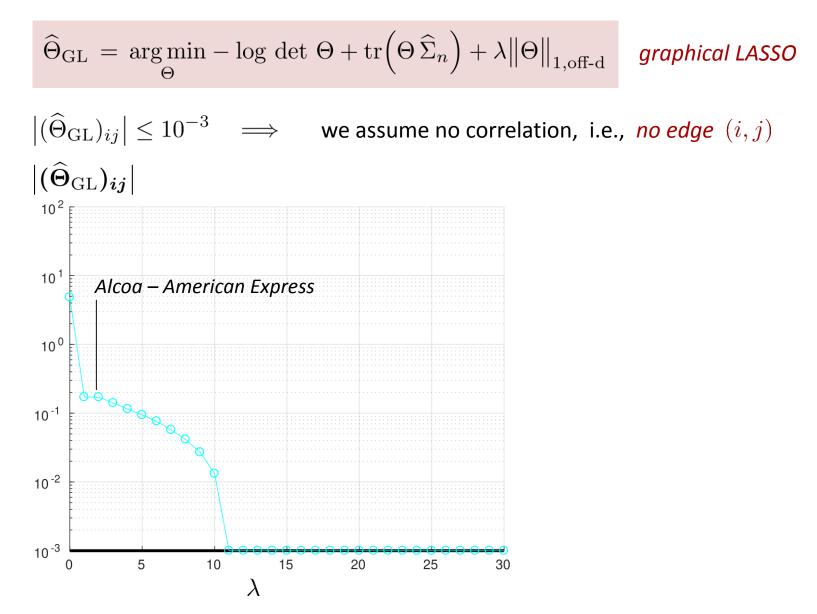
Number of edges



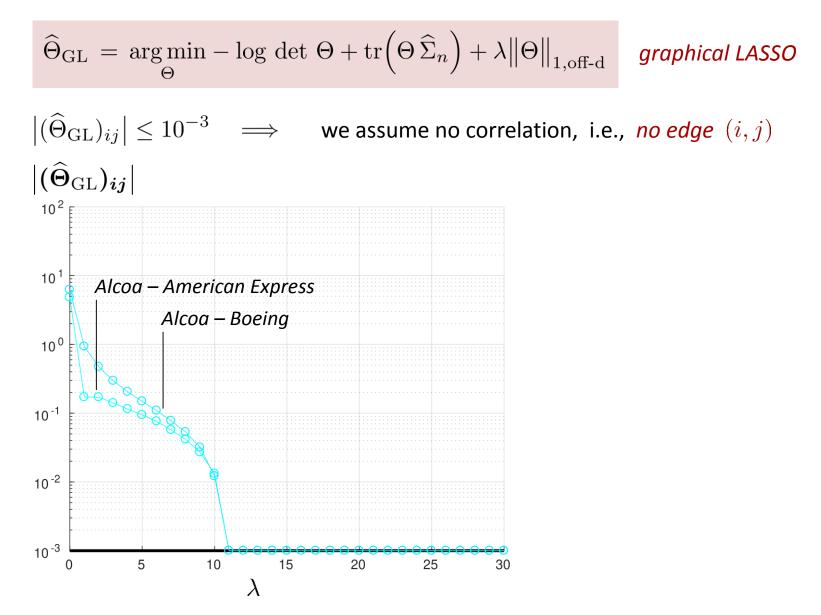




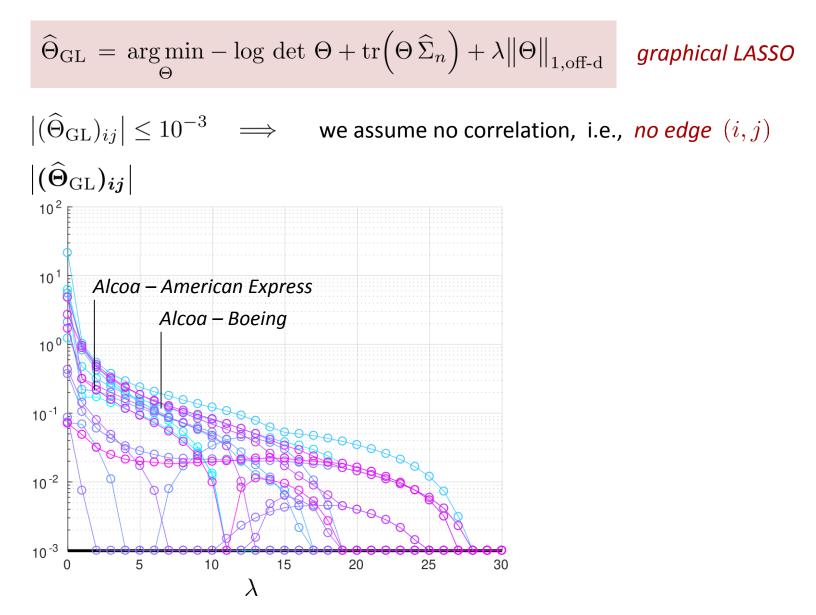




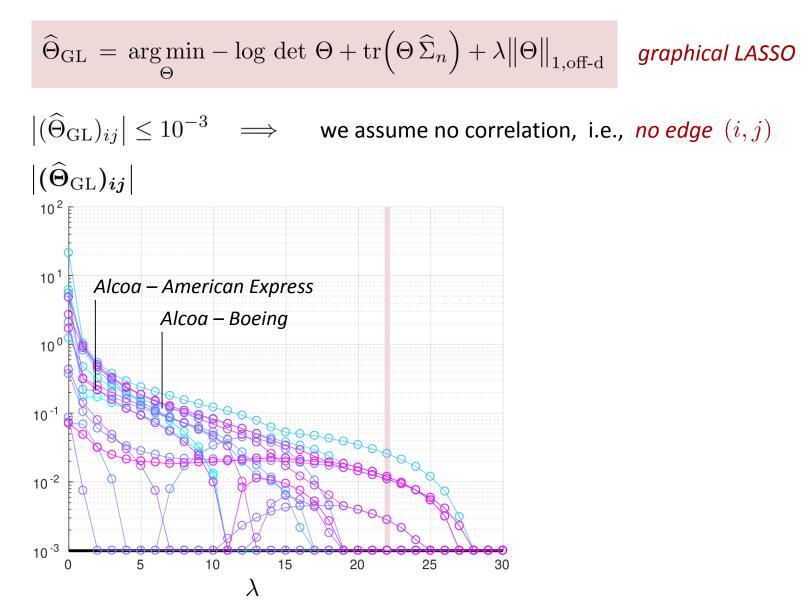




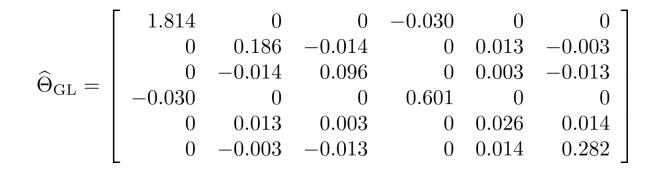




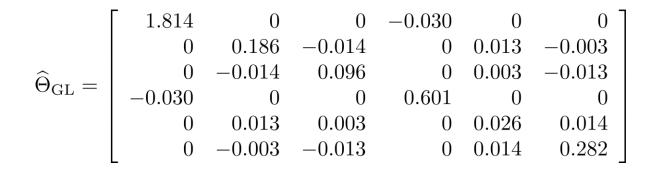












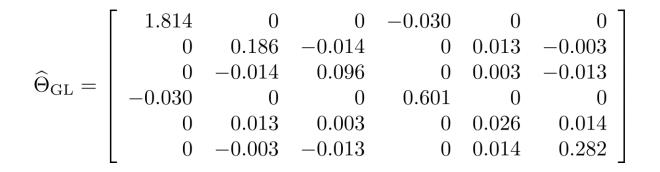
Boeing Bank of America

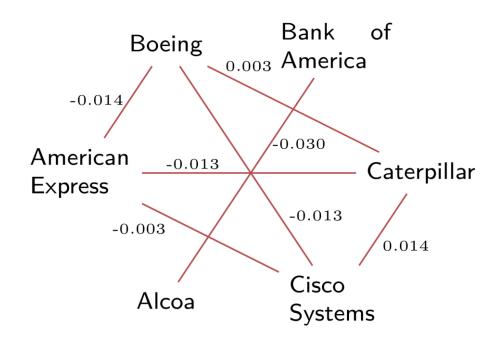
American Express

Caterpillar











Outline

Motivation: Hypothesis Testing in High-Dimensions

Introduction to LASSO and other sparsity problems

Gaussian graphical model selection

Matrix completion





Suppose someone gave you \$1M for *completing* a table like this...

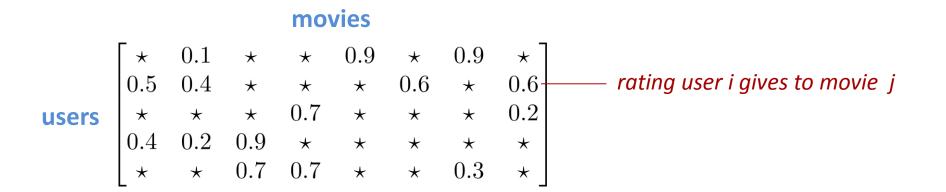


Suppose someone gave you \$1M for *completing* a table like this...

	movies								
users	[*	0.1	*	*	0.9	*	0.9	*]	
	0.5	0.4	*	\star	*	0.6	*	0.6	
	*	\star	*	0.7	\star	*	*	0.2	
	0.4	0.2	0.9	\star	\star	*	\star	*	
	*	*	0.7	0.7	*	*	0.3	*	

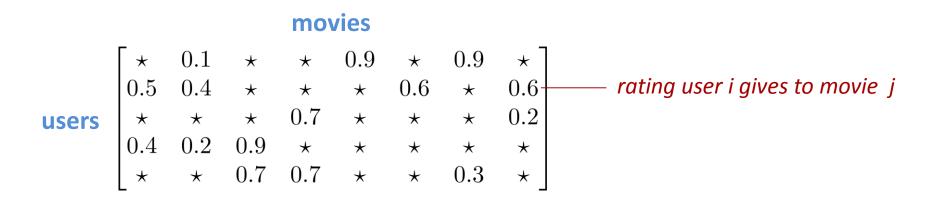


Suppose someone gave you \$1M for *completing* a table like this...





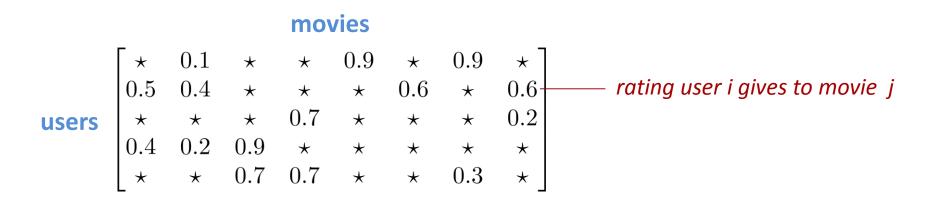
Suppose someone gave you \$1M for *completing* a table like this...



Key insight:

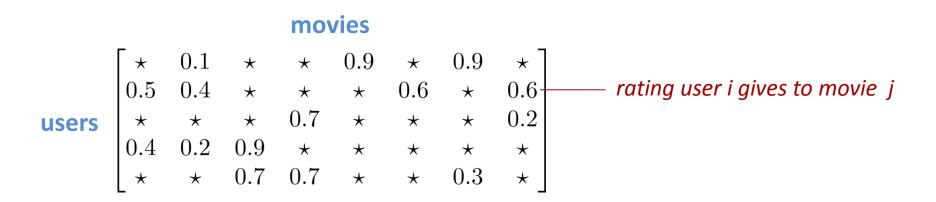


Suppose someone gave you \$1M for *completing* a table like this...





Suppose someone gave you \$1M for *completing* a table like this...

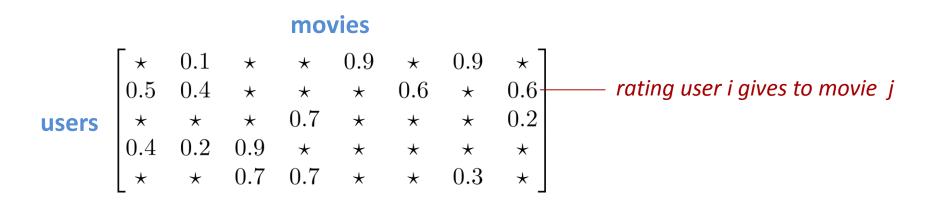


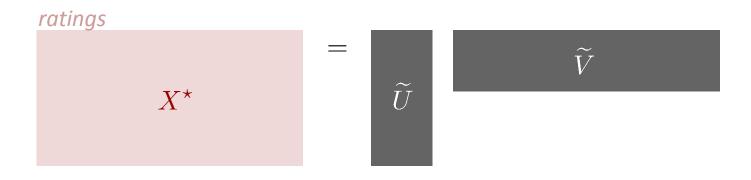
Key insight: only a few factors may explain users' tastes (genre, actors, ads, ...)

ratings



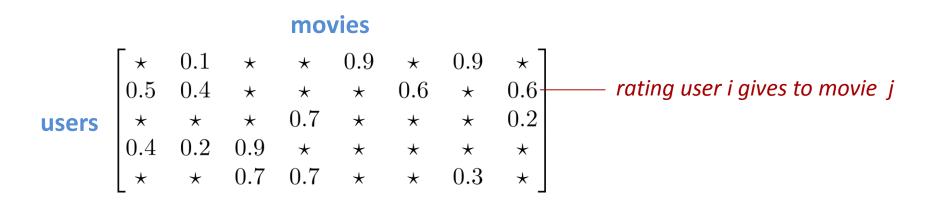
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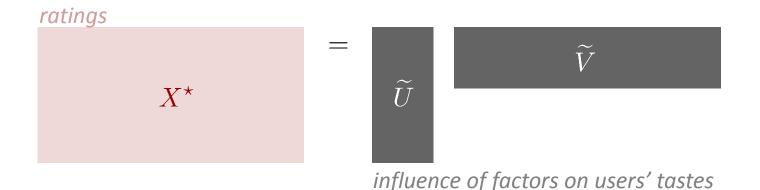






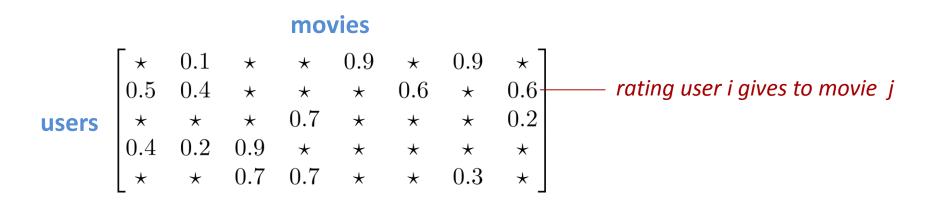
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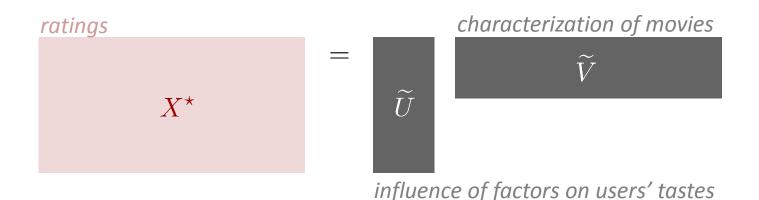






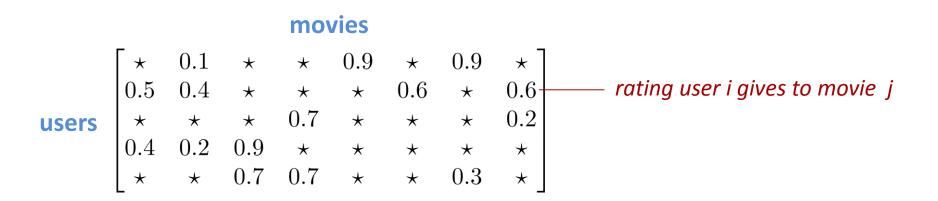
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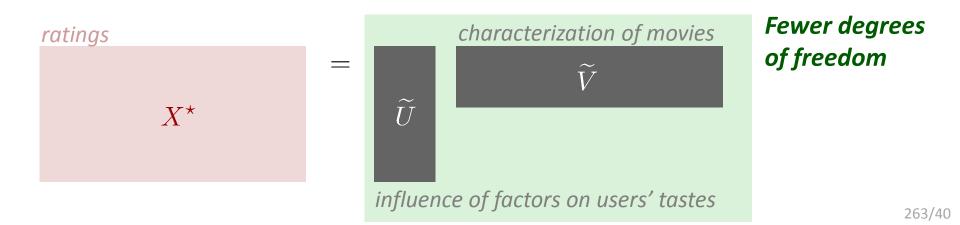






Suppose someone gave you \$1M for *completing* a table like this...







Singular value decomposition:

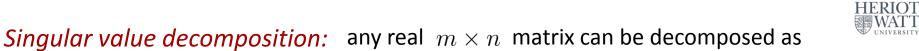


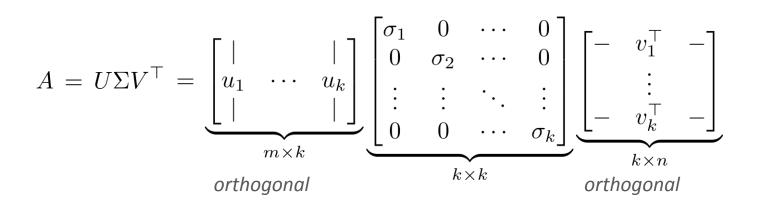


 $A = U\Sigma V^{\top}$



$$A = U\Sigma V^{\top} = \underbrace{\begin{bmatrix} | & & | \\ u_1 & \cdots & u_k \\ | & & | \end{bmatrix}}_{m \times k} \underbrace{\begin{bmatrix} \sigma_1 & 0 & \cdots & 0 \\ 0 & \sigma_2 & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & \sigma_k \end{bmatrix}}_{k \times k} \underbrace{\begin{bmatrix} - & v_1^{\top} & - \\ \vdots & \\ - & v_k^{\top} & - \end{bmatrix}}_{k \times n}$$



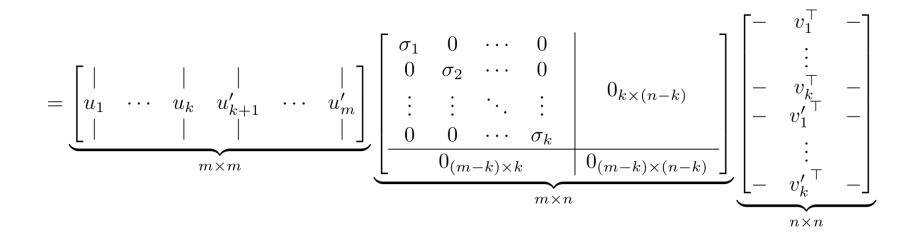




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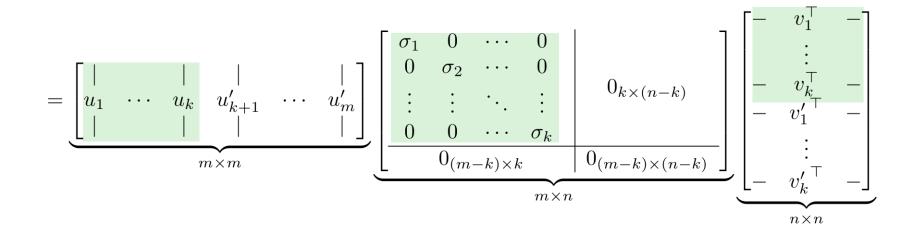


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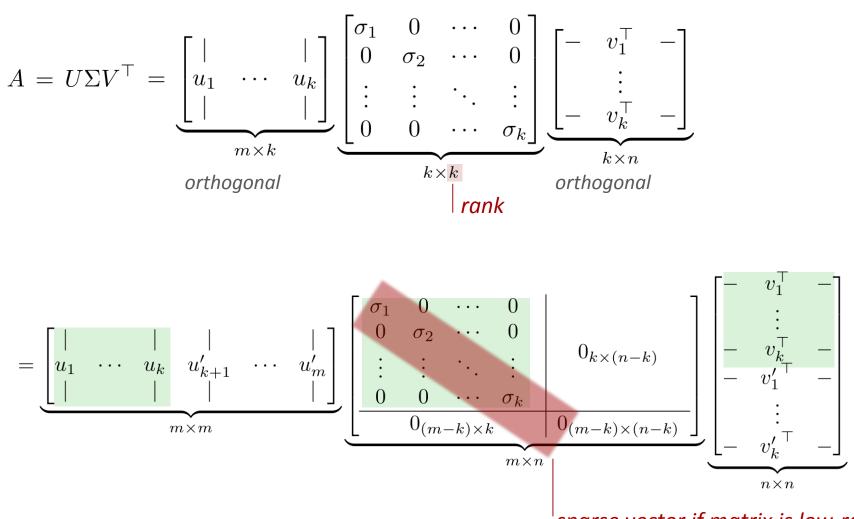




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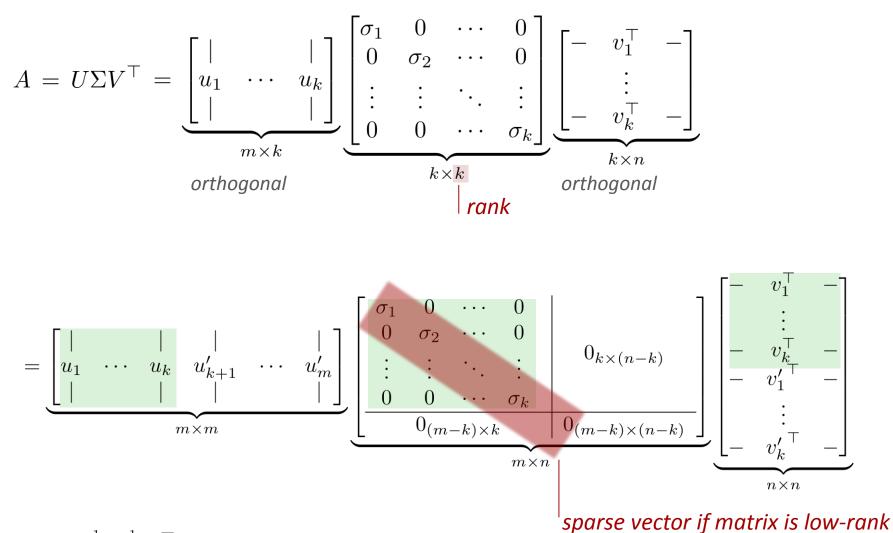






'sparse vector if matrix is low-rank









 $X^{\star} = \underbrace{U\Sigma^{\frac{1}{2}}}_{m \times k} \underbrace{\Sigma^{\frac{1}{2}}V^{\top}}_{k \times n}$



















	movies								
	[*	0.1	*	*	0.9	*	0.9	*]	
users	0.5	0.4	*	*	*	0.6	*	0.6	
	*	\star	\star	0.7	*	*	*	0.2	
	0.4	0.2	0.9	\star	*	*	*	*	
	L *	*	0.7	0.7	*	*	0.3	*	
277/40									







 $\begin{array}{ll} \underset{X \in \mathbb{R}^{m \times n}}{\text{minimize}} & \operatorname{rank}(X) \\ \text{subject to} & X_{ij} = a_{ij} \,, \quad (i,j) \in \mathcal{O} \end{array}$

	movies							
	[*]	0.1	*	*	0.9	*	0.9	*]
users	0.5	0.4	*	*	*	0.6	*	0.6
	*	\star	\star	0.7	\star	*	\star	0.2
	0.4	0.2	0.9	\star	\star	*	\star	*
	*	*	0.7	0.7	*	*	0.3	*
278/40								







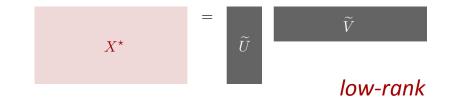
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observed entries

	movies							
	[*	0.1	*	*	0.9	*	0.9	*]
users	0.5	0.4	*	*	*	0.6	*	0.6
	*	\star	\star	0.7	*	*	*	0.2
	0.4	0.2	0.9	\star	*	\star	*	*
	*	*	0.7	0.7	*	*	0.3	*
279/40								







 $\begin{array}{ll} \underset{X \in \mathbb{R}^{m \times n}}{\text{minimize}} & \operatorname{rank}(X) \\ \text{subject to} & X_{ij} = a_{ij} \,, \quad (i,j) \in \mathcal{O} \\ \end{array}$

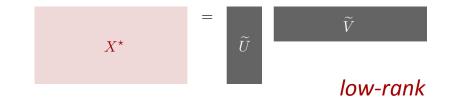
nonconvex

observed entries

	movies							
	[*	0.1	*	*	0.9	*	0.9	*]
users	0.5	0.4	*	*	*	0.6	*	0.6
	*	\star	*	0.7	*	*	*	0.2
	0.4	0.2	0.9	\star	*	*	*	*
	*	*	0.7	0.7	*	*	0.3	*
280/40								







 $\begin{array}{ccc} \underset{X \in \mathbb{R}^{m \times n}}{\text{minimize}} & \operatorname{rank}(X) & & & & & \\ \text{subject to} & X_{ij} = a_{ij} \,, & (i,j) \in \mathcal{O} & & & \\ & & & & & & \\ & & & & & & \\ \text{relax} & & & & \\ \\ \underset{X \in \mathbb{R}^{m \times n}}{\text{minimize}} & \left\| \left(\sigma_1(X), \sigma_2(X), \dots, \sigma_r(X) \right) \right\|_1 \\ \text{subject to} & X_{ij} = a_{ij} \,, & (i,j) \in \mathcal{O} \end{array} \right\|$

 $\begin{array}{c} \text{movies} \\ \text{users} \begin{bmatrix} \star & 0.1 & \star & \star & 0.9 & \star & 0.9 & \star \\ 0.5 & 0.4 & \star & \star & \star & 0.6 & \star & 0.6 \\ \star & \star & \star & 0.7 & \star & \star & \star & 0.2 \\ 0.4 & 0.2 & 0.9 & \star & \star & \star & \star & \star \\ \star & \star & 0.7 & 0.7 & \star & \star & 0.3 & \star \end{bmatrix} \\ 281/40 \end{array}$







 $\begin{array}{ll} \underset{X \in \mathbb{R}^{m \times n}}{\text{minimize}} & \operatorname{rank}(X) & \text{nonconvex} \\ \text{subject to} & X_{ij} = a_{ij} \,, \quad (i,j) \in \mathcal{O} \\ & & & & & & \\ \end{array} \\ \text{relax} \\ \begin{array}{l} \underset{X \in \mathbb{R}^{m \times n}}{\text{minimize}} & \left\| (\sigma_1(X), \sigma_2(X), \dots, \sigma_r(X)) \right\|_1 \\ \text{subject to} & X_{ij} = a_{ij} \,, \quad (i,j) \in \mathcal{O} \end{array} \right\| \\ \end{array}$

	movies								
	[*	0.1	*	*	0.9	*	0.9	*]	
users	0.5	0.4	*	*	*	0.6	*	0.6	
	*	\star	\star	0.7	*	\star	*	0.2	
	0.4	0.2	0.9	*	*	\star	*	*	
	*	*	0.7	0.7	*	*	0.3	*	
282/40									

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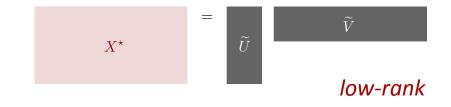
 $\begin{array}{c} \underset{X \in \mathbb{R}^{m \times n}}{\text{subject to}} & \operatorname{rank}(X) & \text{nonconvex} \\ X_{ij} = a_{ij}, & (i,j) \in \mathcal{O} \\ & \text{observed entries} \\ \end{array}$ $\begin{array}{c} \text{relax} & \\ \underset{X \in \mathbb{R}^{m \times n}}{\text{subject to}} & \left\| (\sigma_1(X), \sigma_2(X), \dots, \sigma_r(X)) \right\|_1 \\ \text{subject to} & X_{ij} = a_{ij}, & (i,j) \in \mathcal{O} \end{array} \right\| = \left\| X \right\|_{\star}$

	IIIOVIC5								
	[*	0.1	*	*	0.9	*	0.9	*]	
users	0.5	0.4	*	*	*	0.6	*	0.6	
	*	\star	*	0.7	*	\star	\star	0.2	
	0.4	0.2	0.9	*	*	\star	\star	*	
	L *	*	0.7	0.7	*	*	0.3	*	
283/40									

movies







minimize $\operatorname{rank}(X)$ nonconvex $X \in \mathbb{R}^{m \times n}$ subject to $X_{ij} = a_{ij}, \quad (i,j) \in \mathcal{O}$ observed entries relax nuclear norm $\underset{X \in \mathbb{R}^{m \times n}}{\text{minimize}} \quad \left\| (\sigma_1(X), \sigma_2(X), \dots, \sigma_r(X)) \right\|_1 = \|X\|_{\star}$ subject to $X_{ij} = a_{ij}, \quad (i,j) \in \mathcal{O}$ minimize $||X||_{\star}$ $X \in \mathbb{R}^{m \times n}$

subject to $\operatorname{tr}(XM_l) = a_l$, $l = 1, \ldots, p$

284/40

movies





Theorem [Chandrasekaran et al. 12']



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 $X^{\star} \in \mathbb{R}^{m \times n}$



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 $X^{\star} \in \mathbb{R}^{m \times n}$ unknown, but rank k



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 $p \ge 3k\left(m+n-k\right)+1$



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|^{iid} entries $\mathcal{N}(0,1)$ $a_l = \operatorname{tr}(XM_l), \ l = 1, \dots, p$ measurements

$$p \ge 3k (m + n - k) + 1 \qquad \Longrightarrow \qquad X^* = \underset{X}{\operatorname{argmin}} \quad \|X\|_{\star} \qquad \text{w.h.p.}$$

s.t. $\operatorname{tr}(XM_l) = a_l, \quad l = 1, \dots, p$



 $\begin{array}{ll} \underset{X \in \mathbb{R}^{m \times n}}{\text{minimize}} & \|X\|_{\star} \\ \text{subject to} & \operatorname{tr}(XM_l) = a_l \,, \quad l = 1, \dots, p \end{array}$



 X^{\star} : 30 × 30

$\underset{X \in \mathbb{R}^{m \times n}}{\operatorname{minimize}}$	$ X _{\star}$	
	$\operatorname{tr}(XM_l) = a_l ,$	$l = 1, \ldots, p$



 X^{\star} : 30 × 30

$$\operatorname{rank}(X^{\star}) = 3$$

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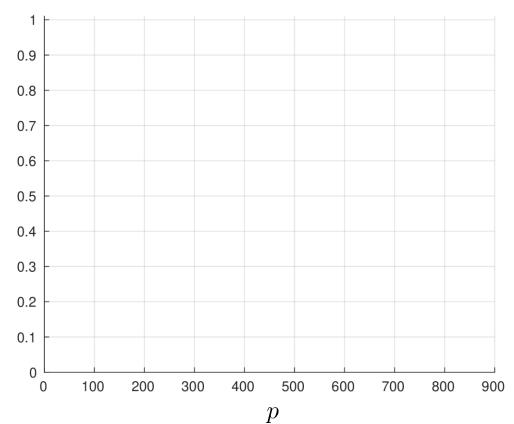


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Sucess rate (20 trials)



297/40

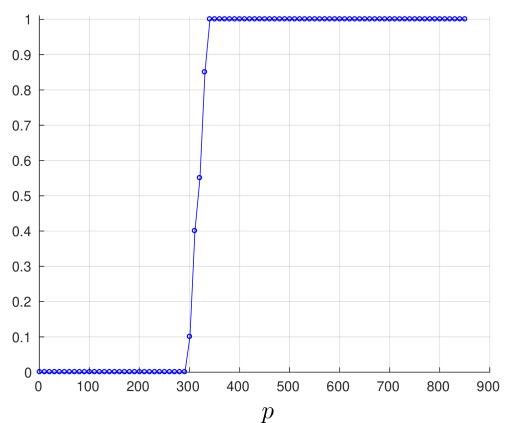


 X^{\star} : 30×30 rank (X^{\star})

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 $||X||_{\star}$ $\underset{X \in \mathbb{R}^{m \times n}}{\text{minimize}}$ subject to $\operatorname{tr}(XM_l) = a_l$, $l = 1, \dots, p$ iid entries $\mathcal{N}(0, 1)$

Sucess rate (20 trials)



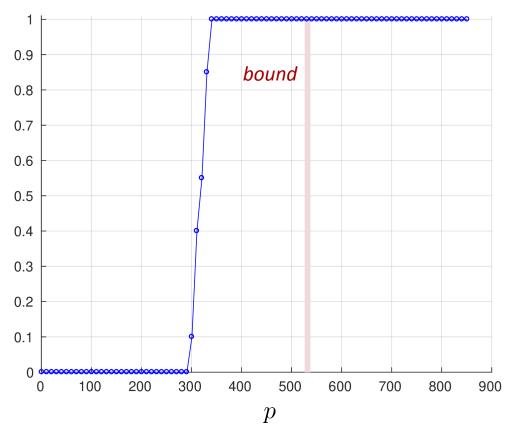


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Sucess rate (20 trials)

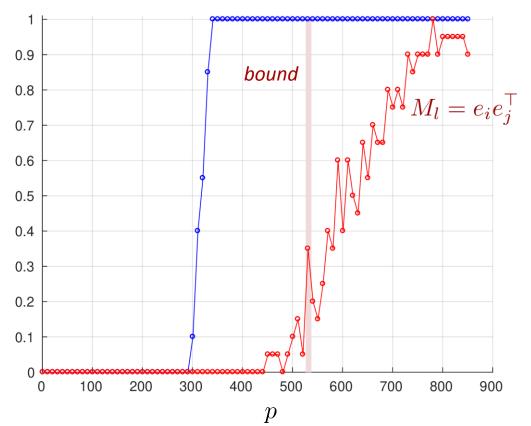




 X^{\star} : 30 × 30

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Sucess rate (20 trials)



300/40

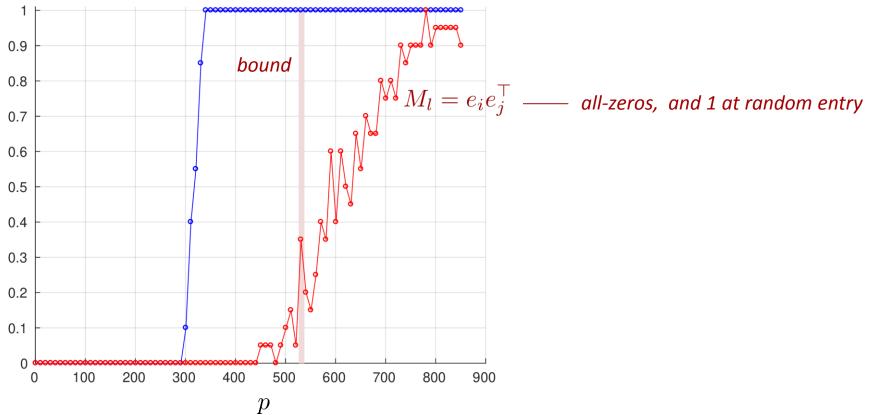


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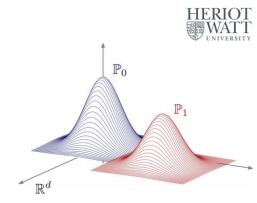
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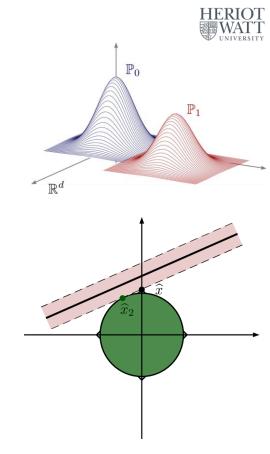




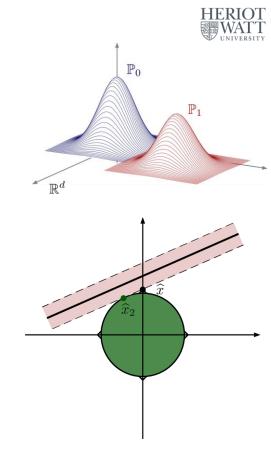
Structure is key in *high-dimensional* problems

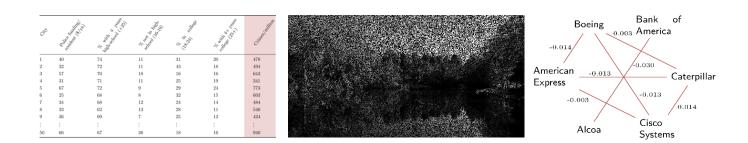


- Structure is key in *high-dimensional* problems
- Sparsity encodes several types of structure

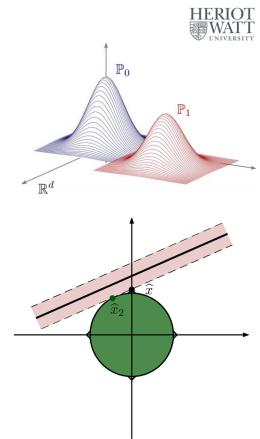


- Structure is key in *high-dimensional* problems
- Sparsity encodes several types of structure
- Several applications (and theory)

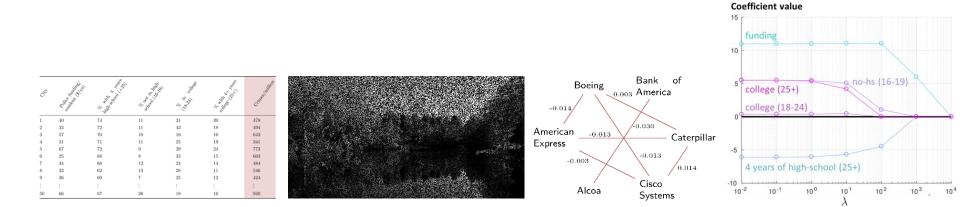




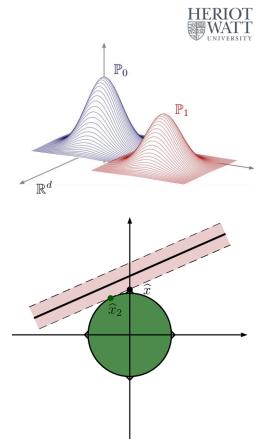
- Structure is key in *high-dimensional* problems
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LASSO, basis pursuit, ... improve *interpretability* and (often) *performance*

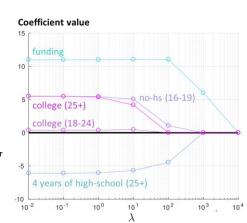


- Structure is key in *high-dimensional* problems
- Sparsity encodes several types of structure
- Several applications (and theory)



- LASSO, basis pursuit, ... improve *interpretability* and (often) *performance*
- Didn't cover: optimization theory and *algorithms*









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V. Chandrasekaran, B. Recht, P. A. Parrilo, A. S. Willsky *The Convex Geometry of Linear Inverse Problems* Foundations of Computational Mathematics, Vol. 12, pp. 805-849, 2012



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M. S. Brown, M. Pelosi, H. Dirska *Dynamic-radius species-conserving genetic algorithm for the financial forecasting of Dow Jones index stocks* Machine Learning and Data Mining in Pattern Recognition, Vol. 7988, pp. 27-41, 2013

Code & presentation

https://github.com/joaofcmota/udrc-summerschool

http://jmota.eps.hw.ac.uk/documents/Mota21-HighDimensionalStatsAndSparsity-UDRC.pdf